

# Generating unfavourable VaR scenarios with patchwork copulas

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# Agenda

- Introduction
- Particular Examples
- A General Framework
- Case Study
- References

## Introduction

Motivation: Create unfavourable scenarios e.g. under Solvency II (i.e. with super-additive VaR, no diversification effects), in particular in the Non-Life segments or for reinsurance purposes (Nat-Cat portfolios)

- Find a “simple” copula construction under which the risk portfolio is in general *unfavourable*
- Give an explicit algorithm for corresponding Monte Carlo studies

## Introduction

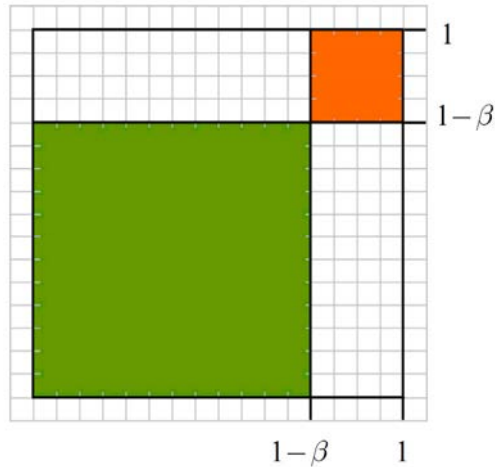
**Lemma 1.** Let, for  $d \geq 2$ ,  $d \in \mathbb{N}$ ,  $\mathbf{U} = (U_1, \dots, U_d)$  and  $\mathbf{V} = (V_1, \dots, V_d)$  be  $d$ -dimensional random vectors over  $[0, 1]^d$  with continuous uniform margins (i.e.,  $\mathbf{U}$  and  $\mathbf{V}$  represent  $d$ -dimensional copulas).

Let further  $I$  denote a binomially distributed random variable, independent of  $\mathbf{U}$  and  $\mathbf{V}$ , with  $P(I = 1) = p = 1 - \beta \in (0, 1)$ . Then the random vector  $\mathbf{W}$  with components

$$W_i := I \cdot p \cdot U_i + (1 - I) \cdot [p + (1 - p) \cdot V_i] \text{ for } 1 \leq i \leq d$$

also has continuous uniform margins, i.e.  $\mathbf{W}$  represents a  $d$ -dimensional copula (patchwork copula, “ordinal sum”).

## Introduction



Shape of the support of the underlying patchwork copula

## Particular Examples

General assumption:  $\mathbf{U}$  and  $\mathbf{V}$  have independent components

**Example 1:** Risks  $X_1$  and  $X_2$  with standard exponential distributions,  
 $S = X_1 + X_2$ ,  $f_S(x, \beta)$  corresponding density

$$f_S(x, \beta) = \begin{cases} \frac{xe^{-x}}{1-\beta}, & 0 \leq x \leq -\ln(\beta) \\ \frac{(-2\ln(\beta) - x)e^{-x}}{1-\beta}, & -\ln(\beta) \leq x \leq -2\ln(\beta) \\ \frac{(x + 2\ln(\beta))e^{-x}}{\beta}, & x \geq -2\ln(\beta) \end{cases}$$

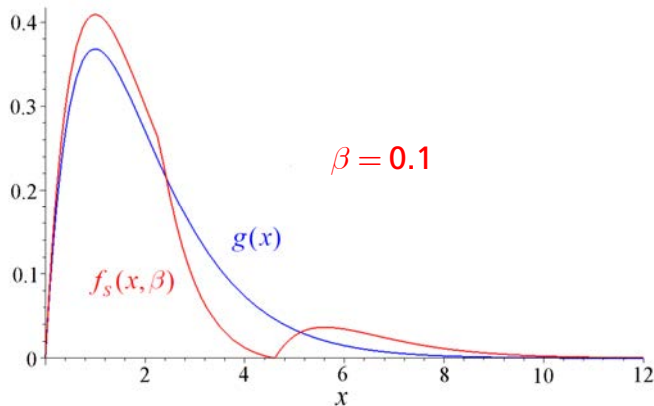
## Particular Examples

**Example 1:** Risks  $X_1$  and  $X_2$  with standard exponential distributions,  $S = X_1 + X_2$ ,  $F_S(x, \beta)$  corresponding cumulative distribution function (cdf)

$$F_S(x, \beta) = \begin{cases} \frac{1 - (1+x)e^{-x}}{1-\beta}, & 0 \leq x \leq -\ln(\beta) \\ \frac{1 - 2\beta + 2e^{-x} \ln(\beta) + (1+x)e^{-x}}{1-\beta}, & -\ln(\beta) \leq x \leq -2\ln(\beta) \\ \frac{\beta - 2e^{-x} \ln(\beta) - (1+x)e^{-x}}{\beta}, & x \geq -2\ln(\beta). \end{cases}$$

## Particular Examples

**Example 1:** Risks  $X_1$  and  $X_2$  with standard exponential distributions,  $S = X_1 + X_2$ ,  $f_S(x, \beta)$  corresponding density,  $g$  Gamma density for independent summands

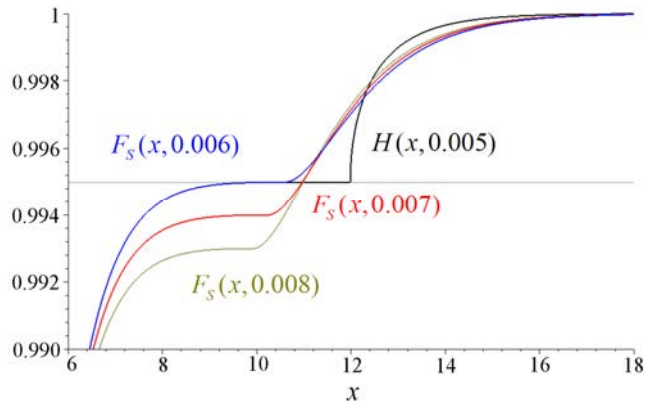




## Particular Examples

**Example 1:** Risks  $X_1$  and  $X_2$  with standard exponential distributions,  $S = X_1 + X_2$ ,  $F_S(x, \beta)$  corresponding cdf,  $H_S(x, \beta)$  worst VaR cdf ( $V =$  lower Fréchet bound)

$\alpha = 0.005$ ;  
Solvency II standard



## Particular Examples

**Example 1:** Risks  $X_1$  and  $X_2$  with standard exponential distributions,  
 $S = X_1 + X_2$  (with patchwork copulas)

| $\alpha = 0.005$                                  | $\beta$ |         |                |         |         |
|---|---------|---------|----------------|---------|---------|
|   | 0.0050  | 0.0060  | 0.0068         | 0.0070  | 0.0080  |
| $\text{VaR}_\alpha(S)$                            | 10.5914 | 10.9630 | <b>10.9829</b> | 10.9821 | 10.9618 |
| $\text{VaR}_\alpha(S)$ (independence)             | 7.4301  | 7.4301  | 7.4301         | 7.4301  | 7.4301  |
| worst $\text{VaR}_\alpha(S)$                      | 11.9829 | 11.9829 | <b>11.9829</b> | 11.9829 | 11.9829 |
| $\text{VaR}_\alpha(X_1) + \text{VaR}_\alpha(X_2)$ | 10.5966 | 10.5966 | 10.5966        | 10.5966 | 10.5966 |

diversification

no diversification

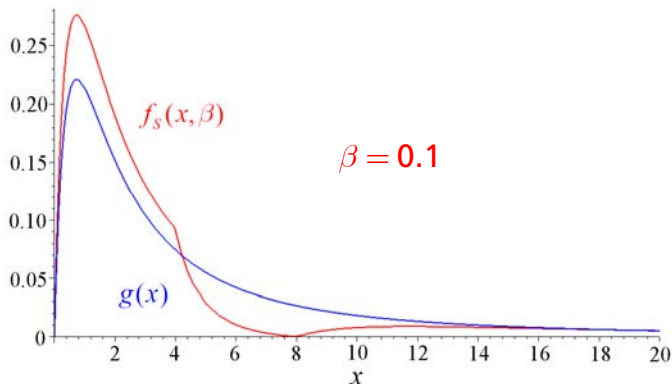
## Particular Examples

**Example 2:** Risks  $X_1$  and  $X_2$  with standard Pareto distributions,  $S = X_1 + X_2$ ,  $F_S(x, \beta)$  corresponding cumulative distribution function (cdf)

$$F_S(x, \beta) = \begin{cases} \frac{x^2 + 2x - 2\ln(1+x)}{(2+x)^2(1-\beta)}, & 0 \leq x \leq \frac{1}{\beta} - 1 \\ \frac{(1-2\beta)x^2 + (4-6\beta)x - 4\beta + 4 + 2\ln(\beta x + 2\beta - 1)}{(2+x)^2(1-\beta)}, & \frac{1}{\beta} - 1 \leq x \leq 2\left(\frac{1}{\beta} - 1\right) \\ \frac{x^2 - 2x + \frac{2}{\beta}\ln(\beta x + 2\beta - 1)}{(2+x)^2}, & x \geq 2\left(\frac{1}{\beta} - 1\right). \end{cases}$$

## Particular Examples

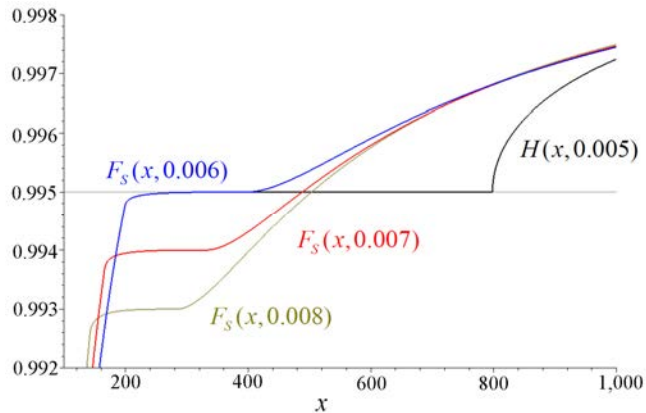
**Example 1:** Risks  $X_1$  and  $X_2$  with standard Pareto distributions,  $S = X_1 + X_2$ ,  $f_S(x, \beta)$  corresponding density,  $g$  density for independent summands



## Particular Examples

**Example 2:** Risks  $X_1$  and  $X_2$  with standard Pareto distributions,  $S = X_1 + X_2$ ,  $F_S(x, \beta)$  corresponding cdf,  $H_S(x, \beta)$  worst VaR cdf ( $V =$  lower Fréchet bound)

$\alpha = 0.005$ ;  
Solvency II standard



## Particular Examples

**Example 2:** Risks  $X_1$  and  $X_2$  with standard Pareto distributions,  
 $S = X_1 + X_2$  (with patchwork copulas)

$\alpha = 0.005$

|   | $\beta$  |          |                 |          |          |
|---|----------|----------|-----------------|----------|----------|
|   | 0.0050   | 0.0070   | 0.0089          | 0.0100   | 0.0110   |
| $\text{VaR}_\alpha(S)$                            | 397.3168 | 503.2848 | <b>509.3804</b> | 508.6489 | 507.0076 |
| $\text{VaR}_\alpha(S)$ (independence)             | 403.9161 | 403.9161 | 403.9161        | 403.9161 | 403.9161 |
| worst $\text{VaR}_\alpha(S)$                      | 798.0000 | 798.0000 | <b>798.0000</b> | 798.0000 | 798.0000 |
| $\text{VaR}_\alpha(X_1) + \text{VaR}_\alpha(X_2)$ | 398.0000 | 398.0000 | 398.0000        | 398.0000 | 398.0000 |

diversification

no diversification

## A General Framework

**Lemma 2.** For  $d \in \mathbb{N}$ ,  $d > 1$  let  $\mathbf{I}_d$  denote the  $d$ -dimensional unit matrix,  $\mathbf{e}_d = (1, \dots, 1)$  the  $d$ -dimensional row vector consisting of just ones, and  $\mathbf{E}_d = \mathbf{e}_d^{\text{tr}} \mathbf{e}_d$  the  $d \times d$  matrix with all entries equal to unity. Then

$$\Sigma_d = (1-r)\mathbf{I}_d + r\mathbf{E}_d = \begin{pmatrix} 1 & r & r & \dots & r \\ r & 1 & r & \dots & r \\ r & r & \ddots & \dots & \vdots \\ \vdots & \vdots & \vdots & 1 & r \\ r & r & \dots & r & 1 \end{pmatrix}$$

is a correlation matrix iff  $-\frac{1}{d-1} \leq r \leq 1$ .

## A General Framework

In the general case, the latent roots  $\lambda_i$  of  $\Sigma_d$  are given by

$$\lambda_1 = 1 + (d-1)r \quad \text{and} \quad \lambda_i = 1 - r, \quad i = 2, \dots, d.$$

An orthonormal basis  $T_1, \dots, T_d$  of corresponding latent vectors is given by

$$T_1 = \frac{1}{\sqrt{d}} \mathbf{e}^{tr} \quad \text{and} \quad T_j = (t_{1j}, \dots, t_{dj})^{tr} \quad \text{for } 2 \leq j \leq d \quad \text{where}$$

$$t_{ij} = \begin{cases} -\frac{1}{\sqrt{j(j-1)}}, & 1 \leq i < j \\ \sqrt{\frac{j-1}{j}}, & j = i \\ 0, & i > j. \end{cases}$$



## A General Framework

Hence  $\Sigma_d$  possesses the spectral decomposition  $\Sigma_d = \mathbf{A}\mathbf{A}^{tr}$  with  $\mathbf{A} = \mathbf{T}\sqrt{\Delta}$  where  $\mathbf{T} = [T_1, \dots, T_d]$  and  $\Delta = \text{diag}(\lambda_1, \dots, \lambda_d)$ .

Note:  $\Sigma_d$  is a particular symmetric Toeplitz matrix for which the latent roots  $\lambda_j^*$  and normalized latent vectors  $T_j^*$  can alternatively be expressed via trigonometric functions, i.e.

$$\lambda_j^* = 1 + r \sum_{i=1}^{d-1} \cos\left(\frac{2\pi ij}{d}\right) = \begin{cases} 1 - r, & j = 1, \dots, d-1 \\ 1 + (d-1)r, & j = d \end{cases}$$

and

$$t_{ij}^* = \frac{\cos\left(\frac{2\pi ij}{d}\right) + \sin\left(\frac{2\pi ij}{d}\right)}{\sqrt{d}}, \quad 1 \leq i, j \leq d.$$

## A General Framework

We call a Gaussian copula derived from the correlation matrix

$$\Sigma_d = \frac{d}{d-1} \mathbf{I}_d - \frac{1}{d-1} \mathbf{E}_d \text{ for } r = -\frac{1}{d-1}$$

a *minimal correlation Gaussian copula*.

Note that the corresponding multivariate normal distribution is degenerated since  $\Sigma_d$  is singular, i.e. a random vector  $\mathbf{X}$  with zero mean and correlation matrix  $\Sigma_d$  has the representation  $\mathbf{X} = \mathbf{A}\mathbf{Y}$  where  $\mathbf{Y}$  has a standard multivariate normal distribution with mean zero and variance-covariance matrix  $\mathbf{I}_d$ . For  $d = 2$ , the minimal correlation Gaussian copula is identical to the lower Fréchet bound or countermonotonicity copula.

## Case Study

The following study shows the effects of such an approach for the 19-dimensional data set discussed in Pfeifer et al. (2019).

It contains insurance losses from a non-life portfolio of natural perils in  $d = 19$  adjacent areas in central Europe over a time period of 20 years.

The losses are given in million monetary units (MMU).

## Case Study

| Year | Area 1  | Area 2  | Area 3  | Area 4 | Area 5 | Area 6 | Area 7  | Area 8 | Area 9 | Area 10 |
|------|---------|---------|---------|--------|--------|--------|---------|--------|--------|---------|
| 1    | 23.664  | 154.664 | 40.569  | 14.504 | 10.468 | 7.464  | 22.202  | 17.682 | 12.395 | 18.551  |
| 2    | 1.080   | 59.545  | 3.297   | 1.344  | 1.859  | 0.477  | 6.107   | 7.196  | 1.436  | 3.720   |
| 3    | 21.731  | 31.049  | 55.973  | 5.816  | 14.869 | 20.771 | 3.580   | 14.509 | 17.175 | 87.307  |
| 4    | 28.99   | 31.052  | 30.328  | 4.709  | 0.717  | 3.530  | 6.032   | 6.512  | 0.682  | 3.115   |
| 5    | 53.616  | 62.027  | 57.639  | 1.804  | 2.073  | 4.361  | 46.018  | 22.612 | 1.581  | 11.179  |
| 6    | 29.95   | 41.722  | 12.964  | 1.127  | 1.063  | 4.873  | 6.571   | 11.966 | 15.676 | 24.263  |
| 7    | 3.474   | 14.429  | 10.869  | 0.945  | 2.198  | 1.484  | 4.547   | 2.556  | 0.456  | 1.137   |
| 8    | 10.020  | 31.283  | 21.116  | 1.663  | 2.153  | 0.932  | 25.163  | 3.222  | 1.581  | 5.477   |
| 9    | 5.816   | 14.804  | 128.072 | 0.523  | 0.324  | 0.477  | 3.049   | 7.791  | 4.079  | 7.002   |
| 10   | 170.725 | 576.767 | 108.361 | 41.599 | 20.253 | 35.412 | 126.698 | 71.079 | 21.762 | 64.582  |
| 11   | 21.423  | 50.595  | 4.360   | 0.327  | 1.566  | 64.621 | 5.650   | 1.258  | 0.626  | 3.556   |
| 12   | 6.380   | 28.316  | 3.740   | 0.442  | 0.736  | 0.470  | 3.406   | 7.859  | 0.894  | 3.591   |
| 13   | 124.665 | 33.359  | 14.712  | 0.321  | 0.975  | 2.005  | 3.981   | 4.769  | 2.006  | 1.973   |
| 14   | 20.165  | 49.948  | 17.658  | 0.595  | 0.548  | 29.350 | 6.782   | 4.873  | 2.921  | 6.394   |
| 15   | 78.106  | 41.681  | 13.753  | 0.585  | 0.259  | 0.765  | 7.013   | 9.426  | 2.180  | 3.769   |
| 16   | 11.067  | 444.712 | 365.351 | 99.366 | 8.856  | 28.654 | 10.589  | 13.621 | 9.589  | 19.485  |
| 17   | 6.704   | 81.895  | 14.266  | 0.972  | 0.519  | 0.644  | 8.057   | 18.071 | 5.515  | 13.163  |
| 18   | 15.550  | 277.643 | 26.564  | 0.788  | 0.225  | 1.230  | 26.800  | 64.538 | 2.637  | 80.711  |
| 19   | 10.099  | 18.815  | 9.352   | 2.051  | 1.089  | 6.102  | 2.678   | 4.064  | 2.373  | 2.057   |
| 20   | 8.492   | 138.708 | 46.708  | 3.680  | 1.132  | 1.698  | 165.600 | 7.926  | 2.972  | 5.237   |

Loss data (I)

## Case Study

| Year | Area 11 | Area 12 | Area 13 | Area 14 | Area 15 | Area 16 | Area 17 | Area 18 | Area 19 |
|------|---------|---------|---------|---------|---------|---------|---------|---------|---------|
| 1    | 1.842   | 4.100   | 46.135  | 14.698  | 44.441  | 7.981   | 35.833  | 10.689  | 7.299   |
| 2    | 0.429   | 1.026   | 7.469   | 7.058   | 4.512   | 0.762   | 14.474  | 9.337   | 0.740   |
| 3    | 0.209   | 2.344   | 22.651  | 4.117   | 26.586  | 3.920   | 13.804  | 2.683   | 3.026   |
| 4    | 0.521   | 0.696   | 31.126  | 1.878   | 29.423  | 6.394   | 18.064  | 1.201   | 0.894   |
| 5    | 2.715   | 1.327   | 40.156  | 4.655   | 104.691 | 28.579  | 17.832  | 1.618   | 3.402   |
| 6    | 4.832   | 0.701   | 16.712  | 11.852  | 29.234  | 7.098   | 17.866  | 5.206   | 5.664   |
| 7    | 0.268   | 0.580   | 11.851  | 2.057   | 11.605  | 0.282   | 16.925  | 2.082   | 1.008   |
| 8    | 0.741   | 0.369   | 3.814   | 1.869   | 8.126   | 1.032   | 14.985  | 1.390   | 1.703   |
| 9    | 0.524   | 6.554   | 5.459   | 3.007   | 8.528   | 1.920   | 5.638   | 2.149   | 2.908   |
| 10   | 9.882   | 6.401   | 106.197 | 44.912  | 191.809 | 90.559  | 154.492 | 36.626  | 36.276  |
| 11   | 1.052   | 8.277   | 22.564  | 8.961   | 19.817  | 16.437  | 25.990  | 2.364   | 6.434   |
| 12   | 0.136   | 0.364   | 28.000  | 7.574   | 3.213   | 1.749   | 12.735  | 1.744   | 0.558   |
| 13   | 1.990   | 15.176  | 57.235  | 23.686  | 110.035 | 17.373  | 7.276   | 2.494   | 0.525   |
| 14   | 0.630   | 0.762   | 25.897  | 3.439   | 8.161   | 3.327   | 24.733  | 2.807   | 1.618   |
| 15   | 0.770   | 15.024  | 36.068  | 1.613   | 6.127   | 8.103   | 12.596  | 4.894   | 0.822   |
| 16   | 0.287   | 0.464   | 24.211  | 38.616  | 51.889  | 1.316   | 173.080 | 3.557   | 11.627  |
| 17   | 0.590   | 2.745   | 16.124  | 2.398   | 20.997  | 2.515   | 5.161   | 2.840   | 3.002   |
| 18   | 0.245   | 0.217   | 12.416  | 4.972   | 59.417  | 3.762   | 24.603  | 7.404   | 19.107  |
| 19   | 0.415   | 0.351   | 10.707  | 2.468   | 10.673  | 1.743   | 27.266  | 1.368   | 0.644   |
| 20   | 0.566   | 0.708   | 22.646  | 6.652   | 14.437  | 63.692  | 113.231 | 7.218   | 2.548   |

### Loss data (II)

## Case Study

A statistical analysis of the data shows a good fit to lognormal  $\mathcal{LN}(\mu, \sigma)$ -distributions for the losses per Area  $k$ ,  $k = 1, \dots, 19$ . The parameters  $\mu_k$  and  $\sigma_k$  for Area  $k$  were hence estimated from the log data by calculating means and standard deviations.

| Parameter  | Area 1 | Area 2 | Area 3 | Area 4 | Area 5 | Area 6 | Area 7 | Area 8 | Area 9 | Area 10 |
|------------|--------|--------|--------|--------|--------|--------|--------|--------|--------|---------|
| $\mu_k$    | 2.8063 | 4.0717 | 3.1407 | 0.6375 | 0.3984 | 1.2227 | 2.3210 | 2.2123 | 1.0783 | 2.1055  |
| $\sigma_k$ | 1.2161 | 1.0521 | 1.2110 | 1.5685 | 1.2998 | 1.5987 | 1.1980 | 0.9882 | 1.1445 | 1.2531  |

| Parameter  | Area 11 | Area 12 | Area 13 | Area 14 | Area 15 | Area 16 | Area 17 | Area 18 | Area 19 |
|------------|---------|---------|---------|---------|---------|---------|---------|---------|---------|
| $\mu_k$    | -0.3231 | 0.3815  | 3.0198  | 1.7488  | 3.0409  | 1.5501  | 3.0700  | 1.2444  | 0.9378  |
| $\sigma_k$ | 1.0881  | 1.3353  | 0.8027  | 1.0033  | 1.1221  | 1.4765  | 0.9622  | 0.8577  | 1.2141  |

Distributional parameters for fitted lognormal loss distributions

## Case Study

|     | A1   | A2    | A3    | A4   | A5   | A6   | A7    | A8    | A9    | A10  | A11  | A12   | A13  | A14  | A15  | A16  | A17  | A18  | A19  |
|-----|------|-------|-------|------|------|------|-------|-------|-------|------|------|-------|------|------|------|------|------|------|------|
| A1  | 1    | 0.46  | 0.03  | 0.16 | 0.47 | 0.20 | 0.35  | 0.49  | 0.41  | 0.24 | 0.78 | 0.64  | 0.91 | 0.63 | 0.85 | 0.66 | 0.30 | 0.67 | 0.56 |
| A2  | 0.46 | 1     | 0.64  | 0.78 | 0.67 | 0.36 | 0.51  | 0.76  | 0.57  | 0.51 | 0.58 | -0.04 | 0.59 | 0.84 | 0.68 | 0.58 | 0.87 | 0.77 | 0.90 |
| A3  | 0.03 | 0.64  | 1     | 0.93 | 0.41 | 0.26 | 0.11  | 0.16  | 0.33  | 0.16 | 0.08 | -0.09 | 0.13 | 0.64 | 0.25 | 0.10 | 0.74 | 0.14 | 0.35 |
| A4  | 0.16 | 0.78  | 0.93  | 1    | 0.54 | 0.36 | 0.16  | 0.25  | 0.43  | 0.19 | 0.22 | 0.00  | 0.30 | 0.79 | 0.36 | 0.19 | 0.84 | 0.32 | 0.49 |
| A5  | 0.47 | 0.67  | 0.41  | 0.54 | 1    | 0.41 | 0.35  | 0.51  | 0.84  | 0.63 | 0.59 | 0.02  | 0.64 | 0.67 | 0.59 | 0.50 | 0.58 | 0.71 | 0.67 |
| A6  | 0.20 | 0.36  | 0.26  | 0.36 | 0.41 | 1    | 0.07  | 0.11  | 0.28  | 0.19 | 0.28 | 0.14  | 0.31 | 0.42 | 0.24 | 0.27 | 0.39 | 0.27 | 0.40 |
| A7  | 0.35 | 0.51  | 0.11  | 0.16 | 0.35 | 0.07 | 1     | 0.44  | 0.27  | 0.19 | 0.48 | -0.07 | 0.46 | 0.35 | 0.45 | 0.91 | 0.64 | 0.61 | 0.49 |
| A8  | 0.49 | 0.76  | 0.16  | 0.25 | 0.51 | 0.11 | 0.44  | 1     | 0.50  | 0.75 | 0.61 | -0.03 | 0.54 | 0.47 | 0.71 | 0.53 | 0.40 | 0.75 | 0.90 |
| A9  | 0.41 | 0.57  | 0.33  | 0.43 | 0.84 | 0.28 | 0.27  | 0.50  | 1     | 0.66 | 0.68 | -0.01 | 0.52 | 0.60 | 0.50 | 0.41 | 0.46 | 0.65 | 0.63 |
| A10 | 0.24 | 0.51  | 0.16  | 0.19 | 0.63 | 0.19 | 0.19  | 0.75  | 0.66  | 1    | 0.33 | 0.00  | 0.27 | 0.28 | 0.43 | 0.24 | 0.23 | 0.45 | 0.65 |
| A11 | 0.78 | 0.58  | 0.08  | 0.22 | 0.59 | 0.28 | 0.48  | 0.61  | 0.68  | 0.33 | 1    | 0.19  | 0.79 | 0.65 | 0.80 | 0.73 | 0.43 | 0.84 | 0.74 |
| A12 | 0.64 | -0.04 | -0.09 | 0.00 | 0.02 | 0.14 | -0.07 | -0.03 | -0.01 | 0.00 | 0.19 | 1     | 0.44 | 0.21 | 0.28 | 0.17 | 0.00 | 0.13 | 0.03 |
| A13 | 0.91 | 0.59  | 0.13  | 0.30 | 0.64 | 0.31 | 0.46  | 0.54  | 0.52  | 0.27 | 0.79 | 0.44  | 1    | 0.71 | 0.86 | 0.74 | 0.47 | 0.76 | 0.65 |
| A14 | 0.63 | 0.84  | 0.64  | 0.79 | 0.67 | 0.42 | 0.35  | 0.47  | 0.60  | 0.28 | 0.65 | 0.21  | 0.71 | 1    | 0.74 | 0.54 | 0.79 | 0.68 | 0.72 |
| A15 | 0.85 | 0.68  | 0.25  | 0.36 | 0.59 | 0.24 | 0.45  | 0.71  | 0.50  | 0.43 | 0.80 | 0.28  | 0.86 | 0.74 | 1    | 0.69 | 0.47 | 0.71 | 0.75 |
| A16 | 0.66 | 0.58  | 0.10  | 0.19 | 0.50 | 0.27 | 0.91  | 0.53  | 0.41  | 0.24 | 0.73 | 0.17  | 0.74 | 0.54 | 0.69 | 1    | 0.63 | 0.77 | 0.64 |
| A17 | 0.30 | 0.87  | 0.74  | 0.84 | 0.58 | 0.39 | 0.64  | 0.40  | 0.46  | 0.23 | 0.43 | 0.00  | 0.47 | 0.79 | 0.47 | 0.63 | 1    | 0.59 | 0.64 |
| A18 | 0.67 | 0.77  | 0.14  | 0.32 | 0.71 | 0.27 | 0.61  | 0.75  | 0.65  | 0.45 | 0.84 | 0.13  | 0.76 | 0.68 | 0.71 | 0.77 | 0.59 | 1    | 0.86 |
| A19 | 0.56 | 0.90  | 0.35  | 0.49 | 0.67 | 0.40 | 0.49  | 0.90  | 0.63  | 0.65 | 0.74 | 0.03  | 0.65 | 0.72 | 0.75 | 0.64 | 0.64 | 0.86 | 1    |

Correlations between losses

yellow = high, green = low

## Case Study

|     | A1   | A2    | A3   | A4   | A5   | A6    | A7    | A8   | A9    | A10  | A11  | A12   | A13  | A14  | A15  | A16  | A17  | A18  | A19  |
|-----|------|-------|------|------|------|-------|-------|------|-------|------|------|-------|------|------|------|------|------|------|------|
| A1  | 1    | 0.27  | 0.30 | 0.16 | 0.17 | 0.45  | 0.28  | 0.32 | 0.32  | 0.29 | 0.67 | 0.51  | 0.76 | 0.34 | 0.67 | 0.74 | 0.18 | 0.21 | 0.29 |
| A2  | 0.27 | 1     | 0.48 | 0.66 | 0.39 | 0.37  | 0.71  | 0.69 | 0.52  | 0.64 | 0.30 | -0.02 | 0.45 | 0.66 | 0.58 | 0.45 | 0.73 | 0.74 | 0.78 |
| A3  | 0.30 | 0.48  | 1    | 0.70 | 0.40 | 0.31  | 0.42  | 0.51 | 0.58  | 0.53 | 0.18 | 0.07  | 0.21 | 0.32 | 0.54 | 0.26 | 0.47 | 0.21 | 0.57 |
| A4  | 0.16 | 0.66  | 0.70 | 1    | 0.77 | 0.47  | 0.46  | 0.47 | 0.59  | 0.49 | 0.18 | 0.08  | 0.33 | 0.50 | 0.47 | 0.18 | 0.76 | 0.43 | 0.54 |
| A5  | 0.17 | 0.39  | 0.40 | 0.77 | 1    | 0.59  | 0.30  | 0.20 | 0.49  | 0.39 | 0.28 | 0.08  | 0.35 | 0.56 | 0.44 | 0.16 | 0.55 | 0.36 | 0.41 |
| A6  | 0.45 | 0.37  | 0.31 | 0.47 | 0.59 | 1     | 0.14  | 0.01 | 0.36  | 0.34 | 0.33 | 0.12  | 0.48 | 0.46 | 0.48 | 0.37 | 0.59 | 0.17 | 0.50 |
| A7  | 0.28 | 0.71  | 0.42 | 0.46 | 0.30 | 0.14  | 1     | 0.52 | 0.27  | 0.40 | 0.45 | -0.07 | 0.31 | 0.31 | 0.46 | 0.62 | 0.63 | 0.58 | 0.57 |
| A8  | 0.32 | 0.69  | 0.51 | 0.47 | 0.20 | 0.01  | 0.52  | 1    | 0.64  | 0.81 | 0.27 | -0.02 | 0.38 | 0.35 | 0.56 | 0.35 | 0.28 | 0.62 | 0.63 |
| A9  | 0.32 | 0.52  | 0.58 | 0.59 | 0.49 | 0.36  | 0.27  | 0.64 | 1     | 0.78 | 0.40 | 0.19  | 0.27 | 0.50 | 0.44 | 0.30 | 0.33 | 0.57 | 0.61 |
| A10 | 0.29 | 0.64  | 0.53 | 0.49 | 0.39 | 0.34  | 0.40  | 0.81 | 0.78  | 1    | 0.21 | -0.02 | 0.21 | 0.37 | 0.52 | 0.30 | 0.31 | 0.53 | 0.81 |
| A11 | 0.67 | 0.30  | 0.18 | 0.18 | 0.28 | 0.33  | 0.45  | 0.27 | 0.40  | 0.21 | 1    | 0.47  | 0.49 | 0.45 | 0.60 | 0.67 | 0.20 | 0.45 | 0.39 |
| A12 | 0.51 | -0.02 | 0.07 | 0.08 | 0.12 | -0.07 | -0.02 | 0.19 | -0.02 | 0.47 | 1    | 0.44  | 0.21 | 0.24 | 0.46 | 0.25 | 0.25 | 0.25 | 0.05 |
| A13 | 0.76 | 0.45  | 0.21 | 0.33 | 0.35 | 0.48  | 0.31  | 0.38 | 0.27  | 0.21 | 0.49 | 0.44  | 1    | 0.55 | 0.60 | 0.71 | 0.37 | 0.39 | 0.24 |
| A14 | 0.34 | 0.66  | 0.32 | 0.50 | 0.56 | 0.46  | 0.31  | 0.35 | 0.50  | 0.37 | 0.45 | 0.21  | 0.55 | 1    | 0.59 | 0.43 | 0.57 | 0.58 | 0.53 |
| A15 | 0.67 | 0.58  | 0.54 | 0.47 | 0.44 | 0.48  | 0.46  | 0.56 | 0.44  | 0.52 | 0.60 | 0.24  | 0.60 | 0.59 | 1    | 0.59 | 0.36 | 0.35 | 0.63 |
| A16 | 0.74 | 0.45  | 0.26 | 0.18 | 0.16 | 0.37  | 0.62  | 0.35 | 0.30  | 0.30 | 0.67 | 0.46  | 0.71 | 0.43 | 0.59 | 1    | 0.38 | 0.43 | 0.39 |
| A17 | 0.18 | 0.73  | 0.47 | 0.76 | 0.55 | 0.59  | 0.63  | 0.28 | 0.33  | 0.31 | 0.20 | 0.35  | 0.37 | 0.57 | 0.36 | 0.38 | 1    | 0.52 | 0.56 |
| A18 | 0.21 | 0.74  | 0.21 | 0.43 | 0.36 | 0.17  | 0.58  | 0.62 | 0.57  | 0.53 | 0.45 | 0.25  | 0.39 | 0.58 | 0.35 | 0.43 | 0.52 | 1    | 0.60 |
| A19 | 0.29 | 0.78  | 0.57 | 0.54 | 0.41 | 0.50  | 0.57  | 0.63 | 0.61  | 0.81 | 0.39 | 0.05  | 0.24 | 0.53 | 0.63 | 0.39 | 0.56 | 0.60 | 1    |

Correlations between log losses

yellow = high, green = low



## Case Study

The following graphs for the tail cdf's correspond to a rank based Bernstein copula **U** with a minimal correlation Gaussian copula **V**:

$$p = 1 [F_1(x)] \text{ (independence case); } p = 0.99 [F_2(x)]; p = 0.994 [F_3(x)]$$

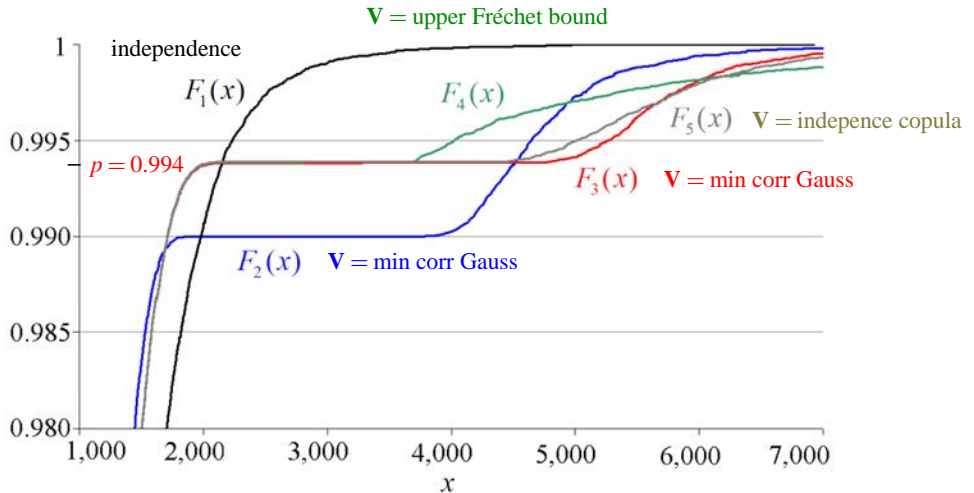
and a rank based Bernstein copula **U** with

$$p = 0.994$$

but different copulas **V**:

upper Fréchet bound  $[F_4(x)]$ ; independence copula  $[F_5(x)]$ .

## Case Study



## Case Study

|  |                |                |               |              |           |
|--|----------------|----------------|---------------|--------------|-----------|
| $\rho$                                   | 0.990          | 0.994          | 0.994         | 0.994        | 1         |
| $\mathbf{V}$                             | min corr Gauss | min corr Gauss | upper Fréchet | independence | ---       |
| $\text{VaR}_\alpha(S)$                   | 4,647 MMU      | 5,272 MMU      | 3,976 MMU     | 5,018 MMU    | 2,229 MMU |
| $\sum_{k=1}^{19} \text{VaR}_\alpha(X_k)$ | 3,976 MMU      | 3,976 MMU      | 3,976 MMU     | 3,976 MMU    | 3,976 MMU |

$$\alpha = 0.005$$

As can clearly be seen, the patchwork construction with the minimal correlation Gaussian copula representing  $\mathbf{V}$  with no tail dependence gives the largest VaR estimate here and is typically larger than for the construction with the upper Fréchet bound which has a positive tail dependence. Note that using the Bernstein copula alone would lead to a truly diversified portfolio while all other copula models do not.

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