

## CONTRIBUTION N° 22

# TRANSACTION COSTS ON THE LONDON TRADED OPTIONS MARKET AND A TEST OF MARKET EFFICIENCY BASED ON A PUT- CALL PARITY THEORY

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COÛTS TRANSACTIONNELS SUR LE  
MARCHÉ DES OPTIONS DE LONDRES  
- UN TEST D'EFFICACITÉ DU MARCHÉ  
FONDÉ SUR LA THÉORIE DE LA  
PARTIE ACHAT-VENTE

# 100 COÛTS TRANSACTIONNELS SUR LE MARCHÉ DES OPTIONS DE LONDRES - UN TEST D'EFFICACITÉ DU MARCHÉ FONDÉ SUR LA THÉORIE DE LA PARITÉ ACHAT-VENTE

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## RESUME

Il est commode de scinder la théorie de la fixation du prix des options en deux volets : d'une part, les modèles d'équilibre général - tels que la théorie de la parité achat - vente avec des conditions limites d'options -, qui sont fondés sur la simple hypothèse que les investisseurs préfèrent gagner plus plutôt que gagner moins ; d'autre part les modèles d'évaluation, dont la validité repose sur des hypothèses distributives sur le cours du titre sous - jacent. Les deux types de modèles dérivent de l'hypothèse que les marchés des actions et des options sont sans viscosité et que le titre sous - jacent ne rapporte aucun dividende pendant la durée de vie de l'option. Des prouesses d'ingéniosité théorique ont été déployées pour tenter de réconcilier le monde réel - où les marchés présentent des viscosités et les actions donnent lieu au paiement de dividendes -, avec l'univers idéal de la théorie des prix des options. La présente étude, qui vise à combler la distance entre la théorie des prix des options, et la réalité empirique des marchés sur lesquels les prix des options sont fixés, a deux objectifs :

- 1 - Evaluer l'impact probable sur une stratégie d'opération fondée sur un modèle, des coûts transactionnels auxquels est confronté un opérateur sur le Marché des options de Londres, et déterminer une méthode solide de prise en compte les coûts transactionnels dans cette stratégie.
- 2 - Tester l'efficacité du Marché des options de Londres en utilisant une théorie de la parité achat - vente, fondée sur les arguments d'arbitrage et donc exempte d'hypothèses sur la dynamique du titre sous - jacent.

La structure des coûts transactionnels sur le Marché des options de Londres est analysé et l'amplitude de l'écart entre l'offre et la demande sur les options, est étudiée en utilisant des évaluations en temps réel des prix au jour le jour sur ce marché. Des comparaisons sont faites entre la distribution des écarts entre l'offre et la demande sur le Marché des options de Londres et celle de la Bourse des options de Chicago, et une méthode améliorée pour tenir compte de l'écart entre l'offre et la demande des options dans des tests d'efficacité du marché est suggérée. Cette méthode de prise en compte des écarts entre l'offre et la demande est utilisée dans un test de l'efficacité du Marché des options de Londres, fondé sur une théorie de la parité achat - vente. Bien que de nombreux écarts à la parité offre - demande aient été constatés, très peu d'entre eux se sont avérés être des occasions de profit exploitables, lorsque l'on tient compte des coûts transactionnels associés à la stratégie d'opération. L'efficacité du marché ne peut donc pas être niée, mais le grand nombre d'écarts à la parité achat - vente infirme l'hypothèse du synchronisme des marchés des actions et des options, sous - jacente aux modèles théoriques d'évaluation des options.

**TRANSACTION COSTS ON THE  
LONDON TRADED OPTIONS MARKET  
AND A TEST OF MARKET EFFICIENCY  
BASED ON PUT-CALL PARITY  
THEORY**

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Option pricing theory can conveniently be split into two groups - **general equilibrium** models, such as option **boundary** conditions and **put-call** parity **theory** which are based **an** the simple assumption that investors prefer more to less, and **valuation models** which rely for their validity on distributional assumptions about the underlying share price. Both types of model were derived on the assumption that the markets for shares and options are frictionless and that the underlying share pays no dividend **during** the life of **the** option. Much subsequent theoretical ingenuity **has** been employed in attempts to reconcile **the** real **world** of market frictions and dividend paying shares with the ideal world of option pricing theory. The current study, **which** attempts to bridge the gap between the theory of option prices and the empirical reality of markets in which option prices are set, has two aims :

1. To assess the likely impact on a model-based trading strategy, of the **transaction** costs facing a trader on the London Traded Options Market and to identify a robust method of **accounting** for transaction costs in the trading strategy.
2. To test the **efficiency** of the London Traded Options Market using **put-call** parity **theory** which is based on arbitrage arguments and is therefore **free** of assumptions about the dynamics of **the** underlying share.

The structure of transaction costs on the London Traded **Options** Market is **analysed** and the size of bid-ask spread on options is studied using real-time **daily price** quotations from the London Traded Options Market. Comparisons between **the** **distribution** of **bid-ask** spreads in London and that on **the** Chicago Board **Options** Exchange are made and an improved method of accounting for option bid-ask spread in market **efficiency** tests is suggested. This method of dealing with bid-ask spreads is used in a test of the **efficiency** of the London Traded Options Market based on put-call parity theory. Although substantial numbers of deviations from put-call parity are identified, very few of these prove to be exploitable profit opportunities when transaction costs associated with the trading **strategy** are accounted for. Although market efficiency cannot therefore be **rejected**, the large number of deviations from put-call parity cast doubt on the assumption of synchronous share and options markets underlying theoretical option valuation models.

## INTRODUCTION

The theory of option pricing can conveniently be split into two groups - general equilibrium models which are based on the simple assumption that investors prefer more to less, and valuation models which rely for their validity on distributional assumptions about the underlying share price. The former category includes boundary conditions on option prices (Merton (1973)), and put-call parity theory which provides a relationship between the prices of puts, calls and the underlying securities (Stoll (1969), Klemkosky and Resnick (1979), Nisbet (1989)). These general equilibrium relationships provide the context for the second group of models, principal among which is the call option valuation model derived by Black and Scholes (1973).

Both types of model were initially derived on the assumption that the markets for shares and options are frictionless and that the underlying share pays no dividend during the life of the option. Much subsequent theoretical ingenuity has been employed in attempts to reconcile the real world of market frictions and dividend paying shares with the ideal world of option pricing theory. The current study, which attempts to bridge the gap between the theory of option prices and the empirical reality of markets in which option prices are set, has two aims :

1. To assess the likely impact on a model-based trading strategy, of the transaction costs facing a trader on the London Traded Options Market and to identify a robust method of accounting for transaction costs in the trading strategy.
2. To employ this method in testing the efficiency of the London Traded Options Market using put-call parity theory. The trading strategy used is based on arbitrage arguments and is therefore free of assumptions about the dynamics of the underlying share.

Previous studies of the London Traded Options Market have been based on the Black-Scholes option pricing model. Both the pricing performance of the Black-Scholes model and the ability of the model to forecast share price volatility have been considered. (Kerruish (1985), Gemmill (1986) and Gemmill and Dickins (1986)). The current study therefore differs from previous studies both by explicitly considering the structure of transaction costs on the LTOM and by testing market efficiency using a test which is free of distributional assumptions about the underlying share price.

Section II of this paper presents the results of a study of the bid-ask spreads facing an arbitrageur on the London Traded Options Market and critically reviews previous methods of accounting for bid-ask spreads in tests of market efficiency. Section III suggests an improved method of accounting for options bid-ask spread and reports the results of using this method in a test of the efficiency of the London Traded Options Market based on put-call parity theory. Section IV presents the conclusions of the paper.

## SECTION II : BID-ASK SPREAD ON THE LONDON TRADED OPTIONS MARKET

### 2.1 Previous Research

An important friction in any options market is the bid-ask spread. The options price spread is the "round trip" costs of **transferring** cash into options and back again, and therefore represents a significant cost of transacting on the options market. Several theories of bid-ask spreads have been suggested.

Neal (1987) develops a model of option bid-ask spreads based on the theory of contestable markets which describes the spread as a function of volume of trading, option price and option volatility, with spread **being** positively correlated with option price and option volatility, and negatively correlated with volume of trading. Neal's model is broadly in agreement in its predictions with models of bid-ask spreads derived by Stoll (1978). Ho and Stoll (1981), Copeland and Galai (1983) and **Glosten** and **Milgrom** (1985). The models by Stoll and Ho and Stoll are based on an examination of the optimal behaviour of a specialist and how this behaviour leads to bid-ask **spreads**. In particular their models predict that, since specialists require compensation for bearing risk, spreads should increase with price volatility. Studies by Benston and **Hagerman** (1974) and by Branch and Freed (1978) find evidence of this type of relationship, but studies by **Tinic** (1972) and **Tinic** and West (1972) are inconclusive on this point. The spread models derived by **Copeland** and Galai and **Glosten** and **Milgrom** relate the spread to informational differences between market participants, with market makers setting the spread to offset losses to informed traders against profits from liquidity traders. **The** model derived predicts that since transaction prices are informative, spreads will tend to decline with volume.

Phillips and Smith (1980). in an empirical study of the bid-ask spread facing traders on the Chicago Board Options Exchange, estimate the cost of liquidity to effect options transactions by calculating the absolute and percentage spreads from **observed** bid and ask quotations in effect on the Chicago Board Options Exchange over four weeks in **1977/1978**. Since there were **restrictions** on trading in options **whose** closing price was less than 50 cents, they report separately options whose prices **are** greater than 50 cents. Their results are presented in Table 2.1.

In the light of their analysis of transaction costs, **Phillips** and Smith reexamine the abnormal profits resulting from options trading strategies reported by Galai (1977, **1978**), **Trippi** (1977). **Chiras** and **Manaster** (1978) and **Klemkosky** and **Resnick** (1979). **Phillips** and **Smith** argue that the most significant transaction cost facing a **trader** is the bid-ask spread and that the abnormal profits reported should therefore be reduced by the bid-ask spreads associated with the relevant trading strategy. They therefore use the **mean** of their observed spread distributions (\$16.05 and 4.51 %) to reduce the profits reported in the above studies and this adjustment effectively eliminates all abnormal profits. It could however be argued that it is inappropriate to use a single figure to adjust **abnormal** profits, since the skewness and dispersion of the **observed** spread distributions reported in Table 2.1 make it unlikely that any single figure could be truly representative.

If for example a **modal** spread of \$6.00, which on the basis of **the** information provided by Phillips and Smith looks like a reasonable estimate, is used **as** the basis of adjustment the abnormal **profits** reported by Galai (1977, 1978) and Klemkosky and Resnick (1979) are not completely eliminated. Conclusions about **market efficiency** base on abnormal profits resulting from a **trading** strategy, are therefore sensitive to **the assumption** made about the appropriate level of spread. Phillips and Smith do not test for possible relationships between spread and the factors, such **as price** and volume of **trading**, identified in the **theoretical** models of bid-ask spreads as having a possible influence on spread

**The** main conclusion of Phillips and Smith is that transaction **costs**, and bid-ask **spreads** in particular, must be accounted for in any test of market **efficiency involving trading rules**. **This** conclusion remains valid. **The** method of adjustment for bid-ask spreads is however critical and whether a single average spread figure applied to **abnormal** profits can **adequately** account for the range of spreads observed by Phillips and Smith remains questionable. The relevance of the Phillips and Smith study to **empirical** studies of the LTOM is dependent on whether or not similar spread distributions to those **observed on the CBOE** can be identified on the London market. Section 2.2. reports the results of a study of the bid-ask spread on the LTOM which attempts to answer this **question**.

## 2.2 Bid-Ask Spread on the **LTOM**

Prior evidence on the size of the bid-ask spread in the London market is largely anecdotal **and** both Gemmill (1986) and Kerruish (1985) refer to large bid-ask **spreads** without formally analysing them, A study was therefore **carried** out to analyse the **spreads** observed **on** the London market. This study **was** a prerequisite to **further work** on option pricing **since** if conclusions about market efficiency were to be obtained by simulating an arbitrageur's trading strategy, it is essential to understand **the nature** and size of transaction **costs** facing the trader on the market. The study reported below therefore provides the framework for the investigations of option market efficiency reported in Section III.

Real time option price quotations obtained by direct feed **from** the floor of **the** London Traded Options **Market** were obtained from **Datstream** for a sample of twelve days during the period 27 June to 22 December 1988. **All shares on** which options were **traded at the** start of the sample period were observed which **resulted** on options on 53 **companies being** included in the sample. Bid and **ask** quotes **for all options** in all classes were obtained for each day in the sample. Two days from each calendar month were chosen and the sample was evenly spread, **as far as possible**, over days of the week. **Table 2.2** gives the dates and corresponding days of the **week** for the sample and **Table 2.3** reports the descriptive statistics for the call and **option** spreads observed. The spread distributions obtained for both calls and puts are given in **Table 2.4**.

Hypothesis **tests** were carried out to determine whether the sample percentage mean spread **of** the calls and puts in **the** British **data** were **significantly different from** those reported by Phillips and Smith for **the** "all options" data. The z values obtained are reported in **Table 2.4** and it can be seen that the **null** hypothesis of no significant

difference between the British and American spreads, is rejected in the case of the puts but is not rejected in the case of the calls at the 1 % level. It would therefore appear that the spreads on calls on the LTOM and the CBOE are similar. However, the medians of the two sets of distributions are different (Calls - LTOM 12.77 % and CBOE 4.23 % ; Puts - LTOM 15.38 % and CBOE 4.76 %) and the relative size of spreads on calls and puts in the two markets is reversed. To investigate further the differences between the observations from the United States and those from the United Kingdom, Kolmogorov-Smirnov one sample and two sample tests were used. Table 2.5 reports the null and alternative hypotheses, the test statistics obtained and the critical values at a 1 % significance level for the Kolmogorov-Smirnov test.

On the basis of these tests, it must be concluded that the spread distributions reported in Phillips and Smith (1980) for the Chicago Board Options Exchange are significantly different from those reported in the current study and that the spreads on the LTOM on both puts and calls are larger than those reported by Phillips and Smith.

Having identified the distribution of bid-ask spreads on the London Traded Options Market, the question then arises of whether it is possible to identify any factors which can be associated with variability in spreads. When market makers on the LTOM were asked how they determined the size of the spread, they unanimously replied that they set the absolute spread in accordance with guidelines drawn up by the Stock Exchange. These guidelines relate the absolute size of the spread to the market price of the option and define the upper limit of the spread to be the size of the spread on the underlying security. *Prima facie*, a relationship between absolute spread and option price could therefore be expected. This relationship could take several forms, the simplest of which is linear of the form :

Absolute Spread =  $a_1 + a_2 \text{ Price}$  where the  $a_i$  are constants.

The model was tested using simple linear regression, with absolute spread as the dependent variable and option mid-market price as the independent variable. Separate regressions were run on the call spreads and on the put spreads and on the calls and puts together. The results of the regressions are contained in Table 2.6. A significant relationship between the price of an option and the absolute bid-ask spread on the option can be seen to exist, with an  $R^2$  value of approximately 0.5 and an F statistic with a significance level of less than 1 %. Therefore option price explains approximately 50 % of the variability in the spread for both puts and calls. In addition, the t statistics obtained imply that both  $a_i$  terms in the regression are significant and the model used therefore looks *prima facie* satisfactory.

It has been argued that, since option series which are far in-the-money or far out-the-money are less likely to be traded than series which are at- or near to-the-money, it is futile to examine prices of option series which are far in-the-money or far out-of-the-money and that spreads on at- or near-the-money options are likely to be more representative than deep in- or deep out-the-money options. [*Quality of Markets Quarterly* (1987)]. The regressions were therefore run again using only option series

which had exercise prices within 10 % of the current share **price**. The results of these regressions are also reported in Table 2.6 and it can be seen that, as before, the relationship of price and spread is **significant** and there is in fact an **improvement** in the explanatory power of option price on spread variability.

### 2.3 Implications for Tests of Market Efficiency

The results **reported** above have significant implications for empirical investigations of options markets based on trading strategies. **Two** principal conclusions can be drawn :

I. The bid-ask spreads on both calls and puts observed on the London Traded Options Market are significantly larger than those reported by Phillips and Smith (1980) for the **Chicago** Board Options Exchange.

The conclusion of Phillips and Smith (1980) that transaction costs, and bid-ask spreads in particular, must be **accounted** for in any test of market efficiency based on trading rules is therefore also relevant to empirical studies of the LTOM.

II. On prices observed on the **London Traded** Options Market, **there** is a statistically **significant** relationship between bid-ask spread and option price.

**This** conclusion implies that the population of bid-ask spreads will change with changes in option prices. If, at a particular time, a sample distribution of bid-ask spreads is identified from the spread **population**, the moments of this distribution **will** reflect the price structure of the spreads contained in the sample. Therefore, it is unlikely that a moment of the sample spread distribution, for example the mean, will be a valid measure of the corresponding moment obtained from a different sample with a **different** price structure. **The** spread adjustment made by Phillips and Smith involves applying the mean of the spread distribution obtained from their sample to abnormal profits derived from samples of options obtained by other researchers at different times. **Although** Phillips and Smith did not attempt to identify explanatory variables for option spread, Neal (1987) regresses option spread against *inter alia* option price and finds a statistically significant relationship. Therefore, because the price structures across samples would almost certainly vary, the sample spread distribution **identified** by Phillips and Smith is unlikely to be representative of the spread distributions faced by the researchers whose findings they examine. The validity of the method of adjustment for **bid-ask** spread used by Phillips and Smith must therefore be questioned on two grounds. Firstly, given the range of spreads observed, it is unlikely that a single figure used for spread adjustment could adequately represent the range of spreads faced by a trader on the markets. Secondly, even if it were possible to **fine** tune the adjustment to reflect the variability in spreads, it is highly unlikely that Phillips and Smith's sample spread distribution could provide a legitimate basis for adjustment in all cases since the price structure of the Phillips and Smith sample is unlikely to be universal. Section **III** reports the results of a test of market efficiency which suggests and uses an alternative method of allowing for bid-ask spread in the trading strategy.

**SECTION III : EMPIRICAL TESTS OF PUT-CALL PARITY ON THE LTOM**

**3.1 The Theory of Put-Call Parity**

Klemkosky and Resnick (1979) derive the following versions of the **put-call** parity theorem which allow for a finite number of known dividends, **d<sub>j</sub>**, **being** paid on the underlying share at **time t<sub>j</sub>**.

$$C_t - P_t - S_t + [K + \sum^n d_j(1+r)^{\delta_j}]/(1+r) \leq 0 \dots (3.1)$$

and  $P_t - C_t + S_t - [K + \sum^n d_j(1+r)^{\delta_j}]/(1+r) \leq 0 \dots (3.2)$

where

- K** = exercise price
- C<sub>t</sub>** = call price at time t
- P<sub>t</sub>** = put price at time t
- S<sub>t</sub>** = share price at time t
- t** = current time
- T** = date of maturity of option (expiry date)
- r** = interest rate over interval [t, T]
- B(t, T-t)** = the price of a **riskless** bond at time t which pays **£1** at time **T**.
- n** = number of dividends paid on share over option period.
- t<sub>j</sub>** = ex-dividend date for **j<sup>th</sup>** dividend
- d<sub>j</sub>** = certain dividend paid at **t<sub>j</sub>**
- τ** = time to **maturity** of option (T-t)
- δ<sub>j</sub>** = the fraction of the option period remaining after the payment of dividend **j**

Inequality 3.1 is derived by considering the cash flows associated with a long hedge **position** consisting of buying a share, writing a call, **buying** a put with the same expiry date and exercise price as the call and borrowing an amount which requires to be repaid on the expiry date by an amount equal to the exercise price. Inequality 3.2 is derived by **considering** the cash flows associated with a short hedge consisting of writing a put, buying a call, shorting the underlying share **and lending** an amount which **will** result in **being** repaid the exercise price on **the** expiry date. **These** inequalities depend for their validity on the assumptions that the arbitrageur is not vulnerable to early exercise of the call in the long **hedge** nor of the put in the short hedge. **Klemkosky** and **Resnick** derive the condition for no early exercise of the call in the long hedge (**inequality** 3.3) and for no rational premature exercise of the put in the short hedge (inequality 3.4).

$$\sum^n d_j(1+r)^{1-\delta_j} < iK/(1+r) \dots \dots \dots (3.3)$$

$$C_t > [iK - \sum^n d_j(1+r)^{\delta_j}]/(1+r) \dots \dots \dots (3.4)$$

Nisbet (1989) extends Klemkosky and Resnick's analysis to take account of transaction costs and derives the following modifications to (3.1) and (3.2).

$$C_t - S_t - P_t + K / (1+r) - T_C - T_S^t - T_P - 2T_S^T / (1+r) + \sum P_j d_j (1+r)^{\delta_j} / (1+r) \leq 0 \dots (3.5)$$

$$P_t - C_t + S_t - K / (1+r) - T_C - T_S^t - T_P - 2T_S^T / (1+r) - \sum P_j d_j (1+r)^{\delta_j} / (1+r) \leq 0 \dots (3.6)$$

where

- $T_C$  = the costs of the call transaction
- $T_S^t$  = the costs of the share transaction at time t
- $T_P$  = the costs of the put transaction
- $T_S^T$  = the costs of the share transaction at time T.

In addition, Nisbet modifies (3.3) and (3.4) to provide the following conditions, in the presence of transaction costs, for no early exercise of the put in the long and short hedges respectively.

$$iK / (1+r) > \sum P_j d_j (1+r)^{\delta_j} / (1+r) - 2T_C - 2T_S^t - T_P - 2T_S^T / (1+r) \dots \dots \dots (3.7)$$

$$C_t > iK / (1+r) - T_C - 2T_S^t - 2T_P - 2T_S^T / (1+r) - \sum P_j d_j (1+r)^{\delta_j} / (1+r) \dots \dots \dots (3.8)$$

**3.2 Data**

The data used in the study were real time option price quotations with synchronous share prices obtained by direct feed from the floor of the London Traded Options Market. Both bid and ask option prices were available and it was therefore possible to explicitly account for bid-ask spread in each hedge by using the bid (ask) price when an option was written (bought). Daily option price quotations and prices for the underlying shares were obtained from Datastream for the period 27 June to 22 December 1988. All shares on which options were traded at the start of the sample period were observed ; consequently options on 53 companies were included in the sample. Since options traded on the LTOM are not dividend payout protected it was necessary to include dividend payments in the model tested. Expected dividend dates and rates were therefore obtained from Datastream for each company included in the sample.

The risk-free rate was estimated by the Inter Bank Rate (IBR) over the relevant period. In an attempt closely to match the interest rate structure facing an arbitrageur, annualised bid and ask rates for one, three, six and twelve months were obtained for each day in the sample period. The interest rate used for each hedge was therefore the annualised bid or ask rate (depending on whether the hedge required borrowing or lending) most closely matched to the period to expiry of the options in the hedge. To obtain an interest rate over the life of the hedge, the annualised rate was prorated to the number of days to expiry of the option, on the assumption that, once borrowing or lending had taken place at that rate, the rate was fixed over the option period

### 3.3 Methodology

The current study employs the standard & vice of establishing hedge portfolios which in equilibrium should return the risk-free rate. Any hedges which return more than the risk-free rate therefore provide a profit opportunity. However the current methodology differs from previous tests of put-call parity in two substantive ways - firstly in the treatment of transaction costs and secondly in testing for a premature termination of the short hedge.

Each day hedges were created for every call price quotation for which there was a corresponding put price quotation, where "corresponding" means that the put is written on the same share and has the same expiration date and exercise price as the call. Any call or put which had a bid or ask price equal to zero was excluded since any such price was necessarily spurious. Over the six months period 27.263 call prices were considered and the hedges created were tested using each of the tests described below. Because of the marginal profits involved in many of the hedges, for each company only the most profitable hedge in any day was used.

Hedges were first created using the relationships between the prices of the call, put, share and exercise price given by (3.1) and (3.2) above. The left hand sides of expressions (3.1) and (3.2) represent the costs of setting up a long and short hedge respectively. If arbitrage is to be avoided and if the hedge position is not to be prematurely terminated by the early exercise of the written call or put, then the upper bound for both expression (3.1) and (3.2) must be zero. If significant numbers of violations of these boundaries are found, put-call parity theory would be rejected and the efficiency of the market would be in doubt.

The mid-market prices of both calls and puts were first used as option price inputs to this model and no account was therefore taken of any transaction costs in these tests, subsequently referred to as tests of Model I, and correspond to the tests carried out by Klemkosky and Resnick (1979) which ignore transaction costs in the model tested.

As explained in Section II, transaction costs have been accounted for in previous tests of put-call parity, not by taking explicit account of transaction costs in the model tested but by identifying a single "reasonable" round trip cost and applying it to all apparently profitable hedges. As has been discussed above, bid-ask spreads show considerable variability and in addition commission rates in the London market are sensitive to the size of the individual transaction. Therefore, tests which apply a flat rate transaction cost to all hedges do not adequately capture the impact of transaction cost variability. In the London Traded Options Market, where both the level and variability of transaction costs are higher than those observed in United States markets, the impact of transaction costs is likely to be even more significant.

In an attempt to isolate the effects of the different elements of transaction costs faced by an arbitrageur, three further models were tested. Model II tests differ from Model I tests only in that they apply option bid price for the written option and the option ask price for the purchased option in the creation of the hedges implied by inequalities (3.1) and (3.2). For example, in the long hedge the call option bid price and put option ask price were

used. Model II tests therefore **account** for the transaction costs associated with option bid-ask spreads, although no account is taken of commission on options, or of transaction costs associated with the share transactions.

Under Models I and II, any hedge which was prima facie profitable was tested for susceptibility to early exercise using inequality (3.3) for the long hedge and inequality (3.4) for the **short** hedge. A short hedge could satisfy inequality (3.4) at inception but still violate it before the expiration date of the **options**. No short hedge is therefore truly risk-free. To obtain an insight into the scale of this problem, **condition (3.4)** was retested each day during the life of each profitable short hedge under both Model I and II and any hedge which violated the condition, and which was therefore vulnerable to early **termination**, was excluded from the set of profitable hedges. This **test**, while obviously having **no** predictive relevance to an arbitrageur, does provide an insight into the scale of the early exercise risk associated with the short hedge portfolio.

It can validly be argued that using option price quotations in the model does not allow actual profit opportunities to be identified, since the price relationships within which profit opportunities can be **identified** are likely to be materially **different** from the prices prevailing at the time transactions actually take place. In an attempt to overcome this problem, **ex-ante** tests were carried out. Any hedge identified as providing a profit opportunity was reestablished one day later using prices then prevailing, **and** was tested both to determine whether or not the hedge had remained profitable and whether it had **become** vulnerable to early exercise over the intervening period.

To take account of remaining transaction costs associated with the hedge strategy, the tests **described** above were repeated using expressions (3.5) and (3.6). **The** transaction cost **terms** in expressions (3.5) and (3.6) derive both from commission charges and **bid-ask** spreads. **As** can be seen **from** Tables 3.1 and 3.2, the level of commission charged on an option transaction depends on the number of contracts included in the bargain and consequently on the value of the transaction, and the share commission rate depends on the total value of the equity transaction. In an attempt to isolate the effects on commission costs of different **transaction** sizes, three different levels of **option** and share **commission** rates were applied to each hedge. These were classified as High, Medium and Low and the following commission rates were used :

Level of Costs	Option Commission	Share Commission
<b>High</b>	6.06%	5.69%
<b>Medium</b>	3.36%	0.86%
<b>Low</b>	0.54%	0.20%

For the underlying shares, however, only the mid-market **prices**, and not the bid and ask quotations were available. Shares are classified on the London International Stock Exchange into Alpha, Beta, Gamma and Delta shares according to how actively they are traded. Alpha share are the most heavily traded and all shares in which options are

traded are Alpha shares. At the end of September 1988, the average touch (the difference between the highest bid price and the lowest ask price over all market makers) for Alpha stocks was 0.80 % and the average spread was 1.45 %. (Quality of Markets Quarterly (Autumn 1988), page 11). To estimate conservatively the spread on shares the higher of these two figures was used and 0.725 % was therefore obtained and applied to the appropriate share price in each hedge.

**All** transaction costs are identifiable at the time the hedge is established excepts  $T_s^T$ , the closing share commission. No hedge created from expressions (2.5) and (2.6) is therefore truly riskless, unless an upper bound can be placed on  $T_s^T$ . Given the structure of transaction costs facing a trader in the London market and the fact that  $T_s^T$  is a function of closing share price, it is not possible to place an upper bound on  $T_s^T$ . In an attempt to assess the impact of the closing transaction costs on the number of profit opportunities, each test was therefore run using two different assumptions. Firstly hedges which allowed for opening transaction costs only were created and tested for profitability using each of the three levels of transaction costs. These are referred to as Model III tests. These tests were then repeated by creating hedges which allowed both for the opening and closing transaction costs. Difficulty is of course encountered in assessing the appropriate level of closing transaction costs. However, if it is assumed that the share price follows a random walk, then the best estimate of the share price at T, the expiration date of the options in the hedge, is the current share price. The commission and spread relevant to the current share price was therefore used as a surrogate for the closing share transaction costs. These tests are referred to as Model IV tests.

Each profitable hedge identified using Models III and IV was then tested for susceptibility to early exercise by using inequalities (3.7) and (3.8) for long hedges and short hedges respectively. For consistency, hedges which were profitable if closing transaction costs were ignored, were tested using expressions (3.7) and (3.8) with closing transaction costs omitted, and hedges which were profitable after full account was taken of transaction costs were tested using the unamended versions of (3.7) and (3.8). Since transaction costs make early exercise less likely, inequalities (3.7) and (3.8) provide a less stringent test for vulnerability to early exercise than inequalities (3.3) and (3.4).

All hedges which were potentially profitable ex-post and were not susceptible to early exercise, were then subjected to the ex-ante test described above.

### 3.5 Results and Interpretation of Put Call Parity Tests

A total of 27,263 call prices were used in creating the hedges for each of Model I and Model II. Each hedge was tested for profitability and vulnerability to premature exercise. Table 3.3 contains a summary of the number of ex-post profitable hedges identified. Consistent with theory and the findings of Klemkosky and Resnick, under both Models there are larger numbers of profitable short hedges than profitable long hedges identified, because of the risk associated with early exercise in the short hedge. Table 3.3. also contains the results of the ex-ante test and it can be seen that even when the ex-ante test is applied to ex-post profitable hedges, substantial numbers of Models I and II

**profit** opportunities remain, indicating prima *facie* violations of **put-call** parity and exploitable **inefficiencies** in the market. **Two** aspects of the **Models I and II** results are interesting. Firstly the effect of accounting for the bid-ask spread in the **Model II** tests is to reduce by half the total number of ex-ante and ex-post profitable hedges. This supports Phillips and Smith (1980) in their conclusion that it is essential to account for the bid-ask spread in any empirical study of market **efficiency**. Secondly, the **surprisingly** large number of profitable hedges identified prima facie contradicts **put-call** parity theory and **casts** doubt on the validity of its assumptions. In particular, the results suggest either the efficiency of the **LTOM** or the equilibrium between the option and share markets must be questioned.

Since it is possible at inception to test whether or not the long hedges are susceptible to **premature termination**, long hedges represent true arbitrage opportunities under Models I and II. Because short hedges may be vulnerable to premature exercise, they are not truly risk-free. Table 3.4 contains, for both Models I and II the results of testing the short hedges which were ex-post profitable for violation of condition (3.4) over the life of the options in the hedge. Under both models, there is a substantial reduction in the number of hedges which remain viable after this test and when these remaining hedges are tested for *ex-ante* profitability there is a further drop in the number which would have provided a **profit** opportunity. Consistent with expectations since the Model I call price is higher, a larger proportion of Model II hedges failed to provide a **profit** opportunity. These results indicate that there is a significant probability of **premature** exercise of the **short** hedged **portfolio** and the **short** hedge in either Model I or Model II cannot therefore be regarded as risk-free. Even the ex-ante short hedge "profit opportunities" in Table 3.3 must therefore be regarded with some **skepticism**.

Table 3.5 shows for both long and short conversions, the numbers of *ex-post* and *ex-ante* Model III profitable hedges (i.e. hedges which remain profitable when the costs of establishing the hedge are reduced from the Model II hedge "profit"). Results are presented for each of the three levels of transaction costs identified above. Consistent with expectations, there is a substantial drop in the numbers of ex-post and ex-ante profitable hedges when the costs of setting up the hedge are accounted for. In addition, the number of hedges which remain profitable after setting-up costs, depends critically on the assumption made about the appropriate level of transaction costs and therefore on the size of the transaction.

When the Model IV profitable hedges are identified by deducting an estimate for the closing share price transaction as described in the Methodology section, from the remaining Model III profit, there is a further substantial reduction in the number of profitable hedges for all levels of **transaction** costs. Table 3.5 summarises the Model IV **findings** and again it can be seen that the number of arbitrage **opportunities** remaining after all transaction costs are accounted for depends critically on the **assumption** made about level of transaction costs.

When interpreting Table 3.5, it must be borne in mind that the "profit opportunities" identified are based on hedges which are not truly **risk-free** since they are subject to **error** in the estimate of the closing share **transaction** costs used. One is therefore drawn to the

**conclusion** that although the Model I and II tests implied large **numbers of** deviations from **put-call** parity, these did not represent exploitable inefficiencies in the market because of the impact of transaction costs. Even an arbitrageur trading in very large amounts (over **£1 million** per transaction) and paying consequently very low **rates** of commission, could only have identified 394 ex-ante profit opportunities out of over 27,000 studied.

#### SECTION IV : CONCLUSIONS

This paper has presented the results of an empirical test of **the** bid-ask spreads **on** the London Traded Options Market and, based **on the** results of that study, has **tested** the efficiency of that market. Four models were tested corresponding to different assumptions about the appropriate level of transaction costs facing an arbitrageur in **the LTOM**.

The empirical results of the Model II test (which ignored transaction **costs** other than bid-ask spread on the options) suggested that there were significant numbers of deviations from **put-call** parity in prices quoted **on the** London **Traded Options** Market. **A large number of the** potentially profitable short conversions became **susceptible** to early termination during their lives and, reflecting this increased risk, a greater number of short hedge than long hedge violations were identified. The Model **III** and IV test results (which measured hedge profitability after all opening transaction costs and all transaction costs respectively) indicate that **these** violations of **put-call** parity are unlikely to represent exploitable inefficiencies in the market.

Previous studies of put-call parity in the **U.S.** have not taken specific account of transaction costs in **the model** tested, but have deducted a single **round trip transaction** cost to emulate the effect of transaction costs on the hedge profit. **The** current study has demonstrated that it is possible to **account** for transaction costs in the **put-call** parity model. Not so to do may result in the identification of spurious profit **opportunities**.

The current study also has implications for research employing option **pricing** models. The evidence presented in this paper can be interpreted in one of two ways. Either it suggests that the equity and options markets are not synchronous or that **there are** inefficiencies (which may not be exploitable after transaction **costs**) in **the** prices quoted on the **LTOM**. If market **synchronisation** and/or efficiency is rejected based **on testing** **put-call** parity, it must also be rejected when a pricing **model** is used **on** the same data base. The current study therefore provides a context for **future** research into **the pricing** structure of the London Traded Options Market.

**Table 2 1 : Distributions of Average Spreads Per Contract Reported by Phillips and Smith (1980)**

% Spreads	CALLS		PUTS	
	All	> 50c.	All	> 50c.
0.0 - 0.2	7	5	7	5
0.2 - 0.4	0	0	0	0
0.4 - 0.6	7	7	3	3
0.6 - 0.8	6	6	0	0
0.8 - 1.0	16	16	3	3
1.0 - 1.2	13	13	6	6
1.2 - 1.5	50	50	6	6
1.5 - 2.0	49	49	19	15
2.0 - 3.0	166	163	109	105
3.0 - 5.0	188	181	181	161
5.0 - 8.0	116	112	132	127
8.0 - 12.0	55	49	70	66
12.0 - 20.0	35	13	44	23
20.0 - 30.0	17	1	11	3
30.0 - 50.0	32	3	9	0
50.0 - 100.0	21	2	6	3
> 100.0	98	0	26	0
n	876	670	632	526
mean (%)	29.85	4.51	15.00	5.77
median (%)	4.23	3.39	4.76	4.44

  

\$ Spreads	CALLS		PUTS	
	All	> 50c.	All	> 50c.
0.00 - 6.25	311	145	169	145
6.25 - 12.50	252	225	220	1%
12.50-25.00	215	213	143	136
25.00-50.00	90	79	83	59
> 50.00	8	8	17	17
mean (\$)	16.05	18.23	18.84	19.10
median (\$)	12.50	12.50	12.50	12.50

**Bble 23 : Days Included in Sample of Bid-Ask Quotations**

Monday	Tuesday	Wednesday	Thursday	Friday
18 July	16 August	28 September	7 October	25 November
14 November	11 October	24 August	6 September	16 December
	20 December		14 July	

**Table 23 : Distributions of Average Spreads Per Contract on the LTOM**

% Spreads	CALLS	PUTS
0.0 - 0.2	0	0
0.2 - 0.4	1	0
0.4 - 0.6	4	0
0.6 - 0.8	3	0
<b>0.9 - 1.0</b>	8	1
1.0 - 1.2	72	26
1.2 - 1.5	0	0
<b>1.5 - 2.0</b>	0	0
2.0 - 3.0	<b>183</b>	<b>97</b>
3.0 - 5.0	680	432
5.0 - 8.0	1397	816
<b>8.0 - 12.0</b>	1753	1385
12.0 - 20.0	1729	<b>1697</b>
20.0 - 30.0	1019	1278
30.0 - 50.0	643	919
50.0 - 100.0	501	837
> 100.0	706	1179
n	8699	8667
mean (%)	29.76	42.69
median (%)	<b>12.77</b>	19.23
standard deviation (%)	46.74	56.59

**Table 2 4 : Descriptive Statistics of **Percentage** Spread Distributions**

	Mean	Standard Deviation	Number of Observations
All option	36.22	52.29	17366
Calls	29.76 (z* = 0.18)	46.74	8699
Puts	42.69 (z* = 45.56)	56.59	8667

z\* is the test statistic from a hypothesis test based on a comparison of the sample mean obtained in the current study with the "all options" mean reported by Phillips and Smith (1980).

**Table 2 5 : Comparison of CBOE and LTOM Spreads-Komogorov-Smirnov Tests**

**Kolmogorov-Smirnov Two Sample Test**

H <sub>0</sub>	:	The population distributions of <b>percentage</b> bid-ask spreads on the <b>LTOM</b> and the CBOE from which the <b>samples</b> of calls ( <b>puts</b> ) were drawn are not <b>significantly</b> different.	
H <sub>1</sub>	:	There is a significant difference between the two distributions.	
		Calls	Puts
Test Statistic		<b>0.464</b>	<b>0.579</b>
Critical value (1 %)		0.043	<b>0.050</b>
Accept/Reject H <sub>0</sub>		Reject	Reject

**Kolmogorov-Smirnov One Sample Test**

H <sub>0</sub>	:	The population distribution of <b>percentage bid-ask</b> spreads on calls ( <b>puts</b> ) on the CBOE is not stochastically larger than the population distribution of <b>percentage</b> bid-ask <b>spreads</b> on calls ( <b>puts</b> ) on the <b>LTOM</b> .	
H <sub>1</sub>	:	There is a <b>significant difference</b> between the two distributions.	
		Calls	Puts
Test Statistic		<b>1476.2</b>	<b>1364.3</b>
Critical value (1 %)		<b>9.2</b>	<b>9.2</b>
Accept/Reject H <sub>0</sub>		Reject	Reject

**Table 2.6 : Simple Linear Regression of Absolute Spreads against Option Price**

<i>Entire Sample</i>						
	$R^2$	F (sig level)	Variable	Regression Parameters		
				$a_i$	Std error	t (sig level)
Calls & Puts (n = 17366)	<b>0.509</b>	17976 (0.00)	Price	0.04	<b>0.0003</b>	134.08 ( <b>0.00</b> )
			Constant	1.89	0.0153	124.19 (0.00)
Calls (n = 8699)	0.521	9477 ( <b>0.00</b> )	Price	0.04	<b>0.0004</b>	97.35 (0.00)
			Constant	1.87	0.0232	80.80 (0.00)
Puts (n = 8667)	<b>0.500</b>	8514 ( <b>0.00</b> )	Price	<b>0.04</b>	<b>0.0005</b>	92.27 ( <b>0.00</b> )
			Constant	1.87	0.0200	93.38 (0.00)
<i>Near-the-money</i>						
	$R^2$	F (sig level)	Variable	Regression Parameters		
				$a_i$	Std error	t (sig level)
Calls & Puts (n = 9397)	0.601	14168 (0.00)	Price	0.07	0.0006	119.03(0.00)
			Constant	1.58	0.0199	79.61 (0.00)
Calls (n = 4703)	<b>0.610</b>	7348 (0.00)	Price	<b>0.07</b>	<b>0.0008</b>	<b>85.72 (0.00)</b>
			Constant	1.48	0.0300	50.10 (0.00)
Puts (n = 4694)	0.624	<b>7802</b> (0.00)	Price	0.08	<b>0.0009</b>	88.34 (0.00)
			Constant	1.56	<b>0.0261</b>	60.23 (0.00)

**Note**

The **significance** levels of the F and t statistics are contained in parenthesis after the values of the test statistics.

**Table 3.1 : Commission Rates on UK Equities - July 1988**

Bargain Size (£)	Average Commission Rate (%)
0 - 600	5.69
601 - 2000	2.00
2001 - 10000	1.46
10001 - 20000	0.86
20001 - 50000	0.47
50001 - 100000	0.31
100001 - 250000	0.25
2500001 - 1 million	0.23
over 1 million	0.20

Source : *Quality & Markets Quarterly*, 1988

**Table 3.2 : Commission Rates on UK Traded Options - July 1988**

Bargain Size (Contracts)	Average Commission Rate (%)
1	6.06
2	2.71
3	3.08
4	2.31
5	2.59
6 - 10	3.36
11 - 50	1.33
51 - 100	1.03
over 101	0.54

Source : *Quality of Markets Quarterly*, 1988**Table 3.3 : Numbers of Ex-Post and Ex-Ante Profitable Hedges From Model I Test**

Hedge Profitable	Model I		Model II	
	<i>ex-post</i>	<i>ex-ante</i>	<i>ex-post</i>	<i>ex-ante</i>
Long Hedges	3602	1729	1090	591
Short Hedges	5477	4789	3093	1955
<b>Total</b>	9079	6518	4183	2546

**Table 3.4 : Numbers of Ex-Post Profitable Short Hedges Not Violating Condition for No Early Exercise Before Option Expiration**

	Model I	Model II
Ex-post profitable short hedges from Table 3.3	5477	3093
Number of ex-post short hedges not violating condition	2235	1012
Number of hedges not violating condition which are profitable ex-ante	1910	478

**Table 35 : Numbers of *Ex-Post* and *Ex-Ante* Profitable Hedges From Model III and Model IV Tests**

Hedge Profitable	MODEL III		MODEL IV	
	<i>ex-post</i>	ex-ante	ex-post	<i>ex-ante</i>
<b>High Transaction costs</b>				
Long Hedge	63	34	40	15
Short Hedge	70	63	20	18
<b>Total</b>	<b>133</b>	<b>97</b>	<b>60</b>	<b>33</b>
<b>Medium Transaction costs</b>				
Long Hedge	235	152	114	60
Short Hedge	819	613	116	99
<b>Total</b>	<b>1054</b>	<b>765</b>	<b>230</b>	<b>159</b>
<b>Low Transaction costs</b>				
Long Hedge	383	205	196	123
Short Hedge	1492	960	423	271
<b>Total</b>	<b>1875</b>	<b>1165</b>	<b>619</b>	<b>394</b>

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