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at ISEG – Lisbon School of Economics
and Management

A matrix approach to pricing marriage insurance with mortality dependence *

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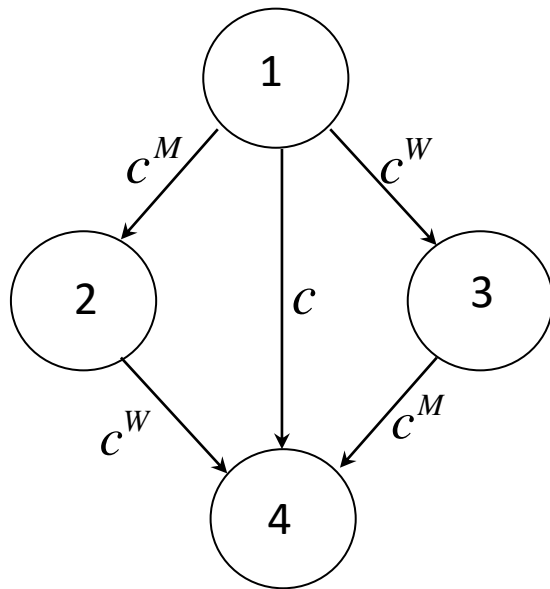
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I. MULTIPLE STATE MODEL FOR MARRIAGE INSURANCE

LSS – Last Surviving Status

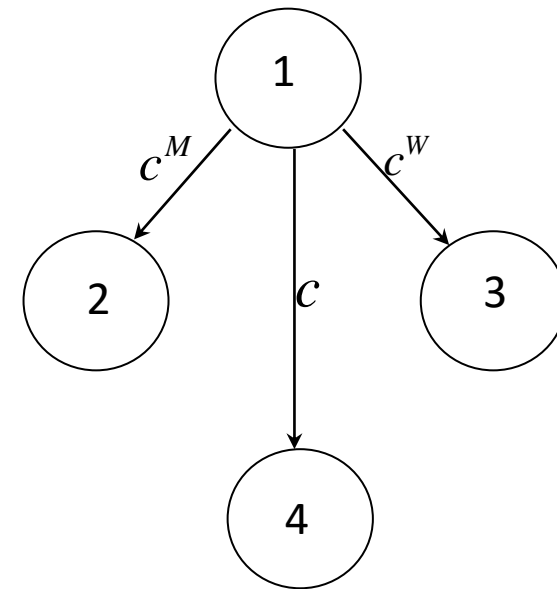


State space:

- 1 - both spouses are alive
- 2 - husband is dead
- 3 - wife is dead
- 4 - both spouses are dead

$$(S, T) = (\{1, 2, 3, 4\}, \{(1, 2), (1, 3), (1, 4), (2, 4), (3, 4)\})$$

JLS – Joint-life Status



$$(S, T) = (\{1, 2, 3, 4\}, \{(1, 2), (1, 3), (1, 4)\})$$



II. PROBABILISTIC STRUCTURE OF THE MODEL

DECREMENTS

x - age at a policy issue of husband

y - age at a policy issue of wife

T_x^M - the remaining lifetimes of x -year-old man (husband); $T_x^M \in [0, w_x^M]$ and $w_x^M = \omega - x$

T_y^W - the remaining lifetimes of y -year-old woman (wife); $T_y^W \in [0, w_y^W]$ and $w_y^W = \omega - y$

n – insurance period

$$n = \max \{w_x^M, w_y^W\} \text{ for LSS}$$

$$n = \min \{w_x^M, w_y^W\} \text{ for JLS}$$

*Maximum
duration of
life according
to Life Tables*

II. PROBABILISTIC STRUCTURE OF THE MODEL

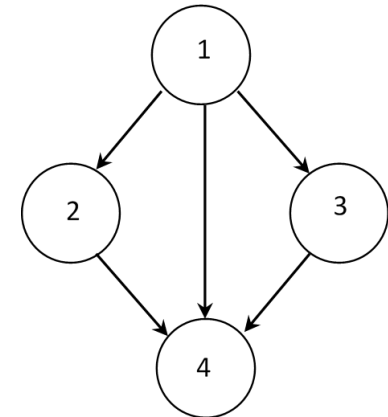
$\{X(t); t \in T\}$ - stochastic process with values in finite set S .

ASSUMPTION: $T = \{0, 1, 2, \dots\}$ - discrete, $\{X(t) : t = 0, 1, 2, \dots\}$ - Markov chain

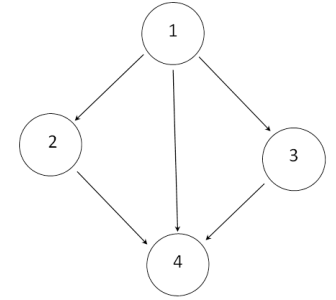
$$\{\mathbf{Q}(k)\}_{k=0}^{n-1}, \text{ where } \mathbf{Q}(k) = (q_{ij}(k))_{i,j=1}^N \text{ and } q_{ij}(k) = P(X(k+1) = j \mid X(k) = i).$$

For marriage insurance with LSS

$$\mathbf{Q}(k) = \begin{pmatrix} q_{11}(k) & q_{12}(k) & q_{13}(k) & q_{14}(k) \\ 0 & q_{22}(k) & 0 & q_{24}(k) \\ 0 & 0 & q_{33}(k) & q_{34}(k) \\ 0 & 0 & 0 & 1 \end{pmatrix}$$



$$T_x^M, T_y^W \text{ - INDEPENDENT } \iff F(w, z) = P(T_x^M \leq w, T_y^W \leq z) = F^M(w) \cdot F^W(z)$$



$$Q(k) = Q^X(k) \circ Q^Y(k)$$

↑
Hadamard (or Schur) product

$$Q^Y(k) = \begin{pmatrix} p_{y+k} & q_{y+k} & p_{y+k} & q_{y+k} \\ 0 & 1 & 0 & 1 \\ 0 & 0 & p_{y+k} & q_{y+k} \\ 0 & 0 & 0 & 1 \end{pmatrix}$$

$$Q^X(k) = \begin{pmatrix} p_{x+k} & p_{x+k} & q_{x+k} & q_{x+k} \\ 0 & p_{x+k} & 0 & q_{x+k} \\ 0 & 0 & 1 & 1 \\ 0 & 0 & 0 & 1 \end{pmatrix}$$

$$T_x^M, T_y^W \text{ - DEPENDENT} \iff F(w, z) = P(T_x^M \leq w, T_y^W \leq z) = C(F^M(w), F^W(z))$$

➤ $F^M(w) = P(T_{x_0}^M \leq w)$ i $F^W(z) = P(T_{y_0}^W \leq z)$

➤ *copula* $C(w, z)$ describes the structure of dependence between $T_{x_0}^M$ and $T_{y_0}^W$

➤ based on the correlation coefficient τ - Kendall for $T_{x_0}^M$ and $T_{y_0}^W$ (men at age $[x_0, \omega)$ and women at age $[y_0, \omega)$) the copula type is selected

➤ if $T_{x_0}^M$ and $T_{y_0}^W$ are independent, then $C(w, z) = w \cdot z$

x_0, y_0 - reference age
for men and women

➤ $S^M(w) = P(T_{x_0}^M > w)$, $S^W(z) = P(T_{y_0}^W > z)$ - *survival functions*

➤ *joint survival function* : $S(w, z) = P(T_{x_0}^M > w, T_{y_0}^W > z)$.

➤ *survival copula* $C^*(w, z)$: $S(w, z) = C^*(S^M(w), S^W(z))$

$$C^*(w, z) = w + z - 1 + C(1 - w, 1 - z)$$



THEOREM

For $(S, T) = (\{1, 2, 3, 4\}, \{(1, 2), (1, 3), (1, 4), (2, 4), (3, 4)\})$ and $x = x_0 + t, y = y_0 + s$ ($t \geq 0$ and $s \geq 0$)

elements of transition matrices $\{\mathbf{Q}(k)\}_{k=0}^{n-1}$ have the following form:

➤ for $k \in \{0, 1, 2, \dots, n-1\}$ and $x+k < \omega$ and $y+k < \omega$

$$q_{11}(k) = \frac{C^*(S_{x_0}^M(t+k+1), S_{y_0}^W(s+k+1))}{C^*(S_{x_0}^M(t+k), S_{y_0}^W(s+k))}$$

$$q_{22}(k) = \frac{C^*(S_{x_0}^M(t), S_{y_0}^W(s+k+1)) - C^*(S_{x_0}^M(t+k), S_{y_0}^W(s+k+1))}{C^*(S_{x_0}^M(t), S_{y_0}^W(s+k)) - C^*(S_{x_0}^M(t+k), S_{y_0}^W(s+k))}$$

$$q_{12}(k) = \frac{C^*(S_{x_0}^M(t+k), S_{y_0}^W(s+k+1)) - C^*(S_{x_0}^M(t+k+1), S_{y_0}^W(s+k+1))}{C^*(S_{x_0}^M(t+k), S_{y_0}^W(s+k))}$$

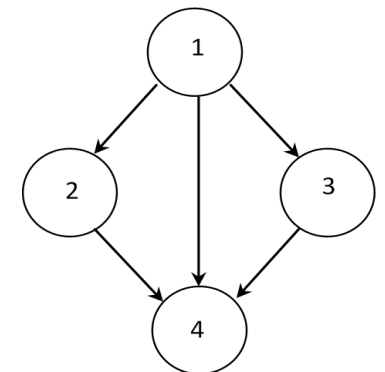
$$q_{24}(k) = 1 - q_{22}(k)$$

$$q_{13}(k) = \frac{C^*(S_{x_0}^M(t+k+1), S_{y_0}^W(s+k)) - C^*(S_{x_0}^M(t+k+1), S_{y_0}^W(s+k+1))}{C^*(S_{x_0}^M(t+k), S_{y_0}^W(s+k))}$$

$$q_{33}(k) = \frac{C^*(S_{x_0}^M(t+k+1), S_{y_0}^W(s)) - C^*(S_{x_0}^M(t+k+1), S_{y_0}^W(s+k))}{C^*(S_{x_0}^M(t+k), S_{y_0}^W(s)) - C^*(S_{x_0}^M(t+k), S_{y_0}^W(s+k))}$$

$$q_{34}(k) = 1 - q_{33}(k)$$

$$q_{14}(k) = 1 - \sum_{j=1}^3 q_{1j}(k)$$





THEOREM

For $(S, T) = (\{1, 2, 3, 4\}, \{(1, 2), (1, 3), (1, 4), (2, 4), (3, 4)\})$ and $x = x_0 + t, y = y_0 + s$ ($t \geq 0$ and $s \geq 0$)

elements of transition matrices $\{Q(k)\}_{k=0}^{n-1}$ have the following form:

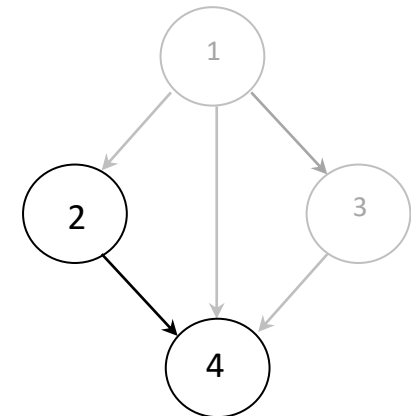
➤ for $k \in \{0, 1, 2, \dots, n-1\}$ and $x+k \geq \omega$ and $y+k < \omega$ (husband exceeded age ω)

$$q_{11}(k) = q_{12}(k) = q_{13}(k) = q_{14}(k) = 0,$$

$$q_{22}(k) = \frac{C^*(S_{x_0}^M(t), S_{y_0}^W(s+k+1)) - C^*(S_{x_0}^M(t+k), S_{y_0}^W(s+k+1))}{C^*(S_{x_0}^M(t), S_{y_0}^W(s+k)) - C^*(S_{x_0}^M(t+k), S_{y_0}^W(s+k))},$$

$$q_{24}(k) = 1 - q_{22}(k),$$

$$q_{33}(k) = q_{34}(k) = 0,$$





THEOREM

For $(S, T) = (\{1, 2, 3, 4\}, \{(1, 2), (1, 3), (1, 4), (2, 4), (3, 4)\})$ and $x = x_0 + t, y = y_0 + s$ ($t \geq 0$ and $s \geq 0$)

elements of transition matrices $\{Q(k)\}_{k=0}^{n-1}$ have the following form:

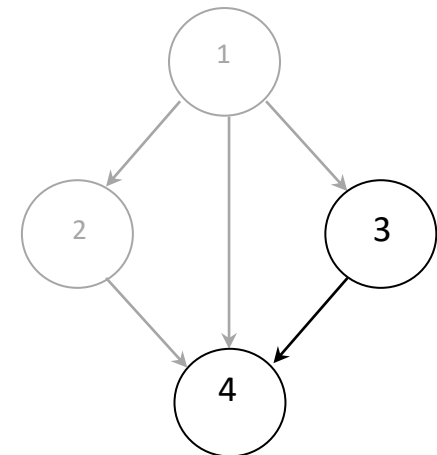
➤ for $k \in \{0, 1, 2, \dots, n-1\}$ and $x+k < \omega$ and $y+k \geq \omega$ (wife exceeded age ω)

$$q_{11}(k) = q_{12}(k) = q_{13}(k) = q_{14}(k) = 0,$$

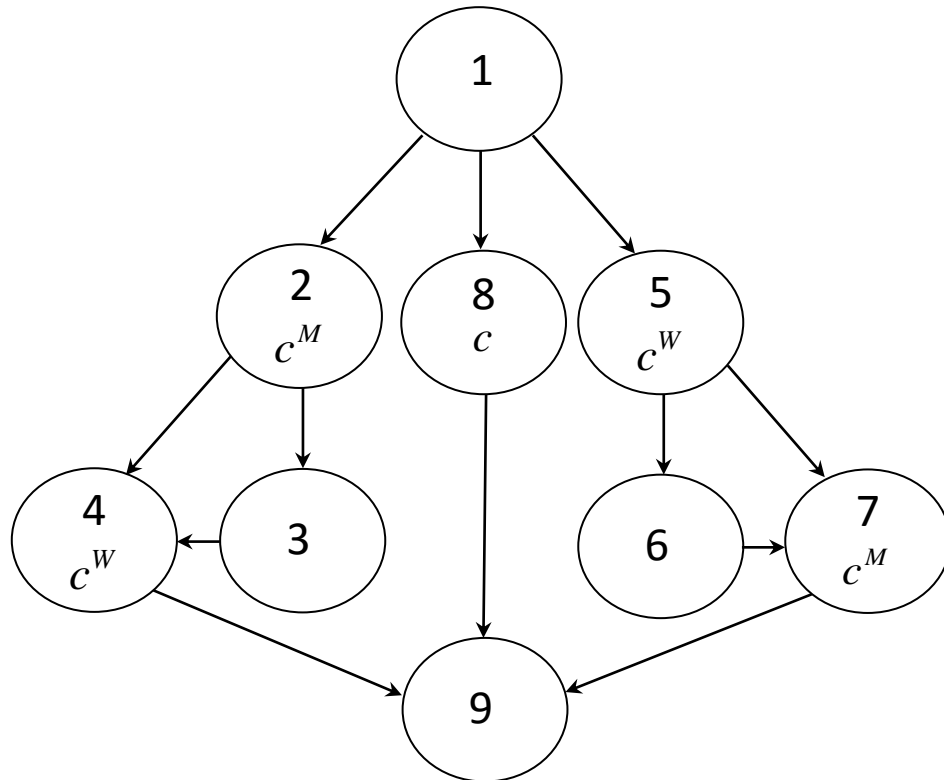
$$q_{22}(k) = q_{24}(k) = 0,$$

$$q_{33}(k) = \frac{C^*(S_{x_0}^M(t+k+1), S_{y_0}^W(s)) - C^*(S_{x_0}^M(t+k+1), S_{y_0}^W(s+k))}{C^*(S_{x_0}^M(t+k), S_{y_0}^W(s)) - C^*(S_{x_0}^M(t+k), S_{y_0}^W(s+k))},$$

$$q_{34}(k) = 1 - q_{33}(k).$$



III. MATRIX FORM OF ACTUARIAL VALUES

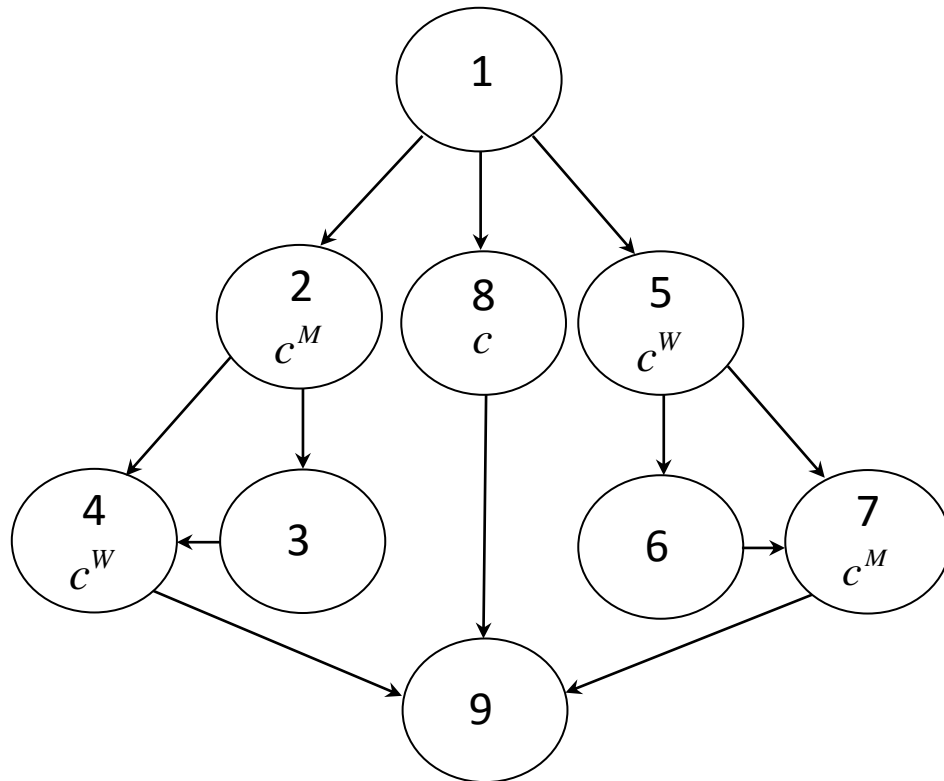


State space of *modified multiple state model* (s^*, T^*)

- 1 - both spouses are alive
- 2 - death of the husband, wife is alive
- 3 - husband has been dead for at least one year, wife is alive
- 4 - husband has been dead for at least one year, death of the wife
- 5 - death of the wife, husband is alive
- 6 - wife has been dead for at least one year, husband is alive
- 7 - wife has been dead for at least one year, death of the husband
- 8 - both spouses are dead
- 9 - both spouses have been dead for at least one

$$Q^*(k) = \begin{pmatrix} q_{11}(k) & q_{12}(k) & 0 & 0 & q_{13}(k) & 0 & 0 & q_{14}(k) & 0 \\ 0 & 0 & q_{22}(k) & q_{24}(k) & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & q_{22}(k) & q_{24}(k) & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 0 & 0 & q_{33}(k) & q_{34}(k) & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & q_{33}(k) & q_{34}(k) & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 \end{pmatrix}$$

III. MATRIX FORM OF ACTUARIAL VALUES



Matricies:

D – matrix describing the probabilistic structure of the model (based on $\{Q^*(k)\}_{k=0}^{n-1} : \mathbf{P}^T(t) = \mathbf{P}^T(0) \prod_{k=0}^{t-1} Q^*(k)$)

C_{in} – benefits matrix

C_{out} – premiums matrix

C = **C_{in}** + **C_{out}** – cash flow matrix

Λ - matrix describing the interest rate (constant, function depending on time or stochastic)

$$\mathbf{S} = (1, 1, \dots, 1)^T \in \mathbb{R}^{N^*}$$

$$\mathbf{J}_i = \left(0, 0, \dots, 0, \underset{i}{1}, 0, \dots, 0 \right)^T \in \mathbb{R}^{N^*}$$

$$\mathbf{I}_{t+1} = \left(0, 0, \dots, 0, \underset{t+1}{1}, 0, \dots, 0 \right)^T \in \mathbb{R}^{n+1}$$



III. MATRIX FORM OF ACTUARIAL VALUES

- **net single premium π** , paid in advance, when $X^*(0) = 1$

$$\pi = \mathbf{S}^T \text{Diag}(\mathbf{C}_{in} \mathbf{D}^T) \mathbf{\Lambda} \mathbf{I}_1$$

- **net period premium p** payable in advance at the beginning of the time unit during the first m units ($1 \leq m \leq n^*$), if $X^*(t) = 1$

$$p = \frac{\mathbf{S}^T \text{Diag}(\mathbf{C}_{in} \mathbf{D}^T) \mathbf{\Lambda} \mathbf{I}_1}{\mathbf{I}_1^T \mathbf{\Lambda}^T [\mathbf{I} - \sum_{t=m+1}^{n+1} \mathbf{I}_t \mathbf{I}_t^T] \mathbf{D} \mathbf{I}_1}$$

- **vector of net prospectives reserves at moment t**

$$\mathbf{V}(t) = \left(\mathbf{C}_{out}^T + \mathbf{C}_{in}^T + \mathbf{F}^T(t, \mathbf{C}) \mathbf{\Lambda} \right) \mathbf{I}_{t+1}$$

where

$$\mathbf{F}^T(t, \mathbf{C}) = \sum_{k=t+1}^n \prod_{u=t}^{k-1} \mathbf{Q}(u) \mathbf{C}^T \mathbf{I}_{k+1} \mathbf{I}_{k+1}^T$$



IV. NUMERICAL EXAMPLES

BENEFITS: $c^M = c^W = 1$; $c = c^M + c^W = 2$

PROBABILISTIC STRUCTURE:

➤ independence: Lower Silesia Life Table 2011

➤ dependence: Main Statistical Office – data for Lower Silesia 2011; $x_0 = y_0 = 60$
Kendall correlation coefficient between T_{60}^M and T_{60}^W : $\tau = 0.076$

Gumbel copula: $C(u, v) = \exp\left(-\left((-\ln u)^\alpha + (-\ln v)^\alpha\right)^{\frac{1}{\alpha}}\right)$ where $\alpha = 1.0786$.

INTEREST RATE:

➤ the real Polish market data related to the yield to maturity on fixed interest bonds from 03.03.2015

➤ the best fitted model is **Svensson model** => term interest rate is equal to **2.1%**

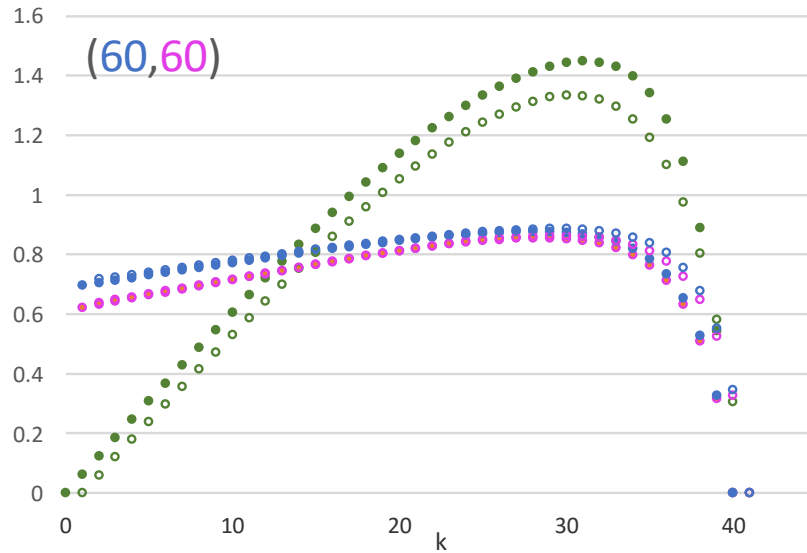


NET PREMIUMS

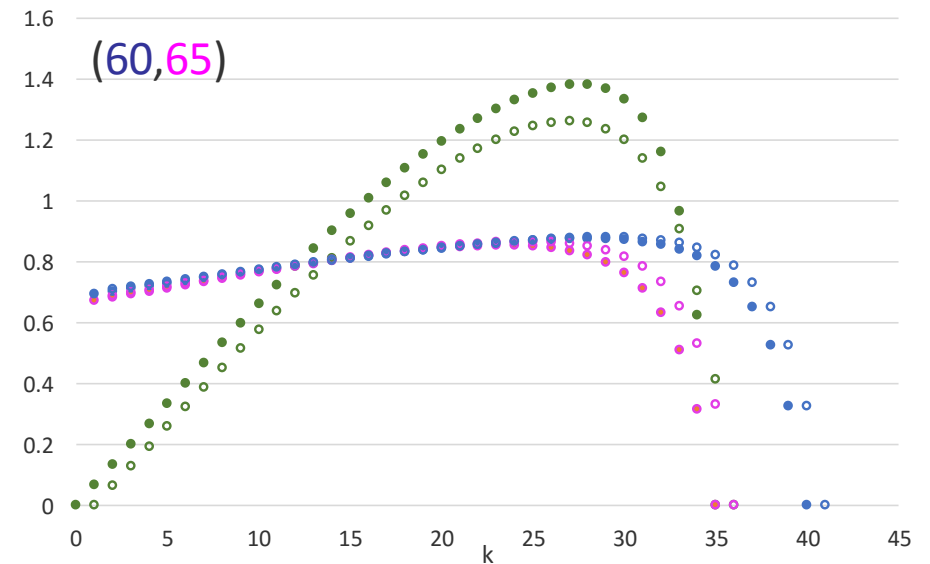
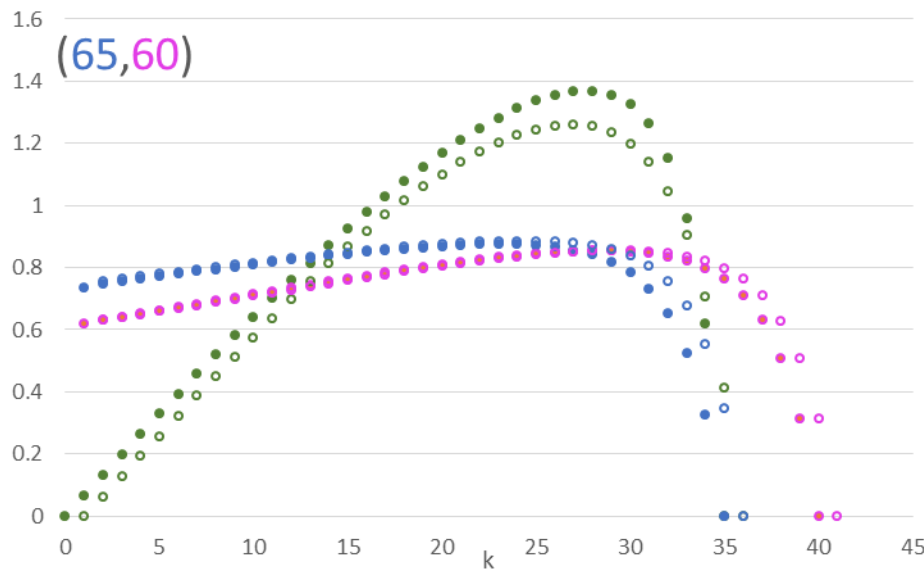
| $(\text{husband age, wife age}) = (x, y)$ | (60, 60) | (65, 60) | (60, 65) | (65, 65) |
|--|----------|----------|----------|----------|
| <i>The insurance period (n)</i> | 40 | 40 | 40 | 35 |
| <i>The period of premiums payment (m)</i> | 40 | 35 | 35 | 35 |
| T_y^M, T_x^W - INDEPENDENT | | | | |
| <i>Net single premium - π</i> | 1.298378 | 1.341719 | 1.348619 | 1.392516 |
| <i>Net period premium - p</i> | 0.100589 | 0.116768 | 0.112917 | 0.128969 |
| T_y^M, T_x^W - DEPENDENT | | | | |
| <i>Net single premium - π</i> | 1.298377 | 1.338764 | 1.346594 | 1.389498 |
| <i>Net period premium - p</i> | 0.099013 | 0.114646 | 0.110538 | 0.125823 |



NET PROSPECTIVE RESERVES



- both spouses are alive (independence)
- wife is alive (independence)
- husband is alive (independence)
- both spouses are alive (dependence)
- wife is alive (dependence)
- husband is alive (dependence)



AVERAGE PROFIT EXPECTED TO EMERGE AT THE END OF YEAR k

$$\mathbf{EPR} = \begin{pmatrix} EPR(1) \\ EPR(2) \\ \vdots \\ EPR(n) \end{pmatrix}$$

$$EPR(k) = \mathbf{I}_k^T \cdot \mathbf{D} \cdot \left(\left(-\mathbf{C}_{out}^T \mathbf{I}_k \mathbf{I}_k^T \tilde{\Lambda} - \mathbf{Q}(k-1) \mathbf{C}_{in}^T \right) \mathbf{I}_{k+1} + \left(\mathbf{V}^T \mathbf{I}_k \mathbf{I}_k^T \tilde{\Lambda} - \mathbf{Q}(k-1) \mathbf{V}^T \right) \mathbf{I}_{k+1} \right).$$

FIRST-ORDER BASIS (prudent)

T_x^M, T_x^W - independent

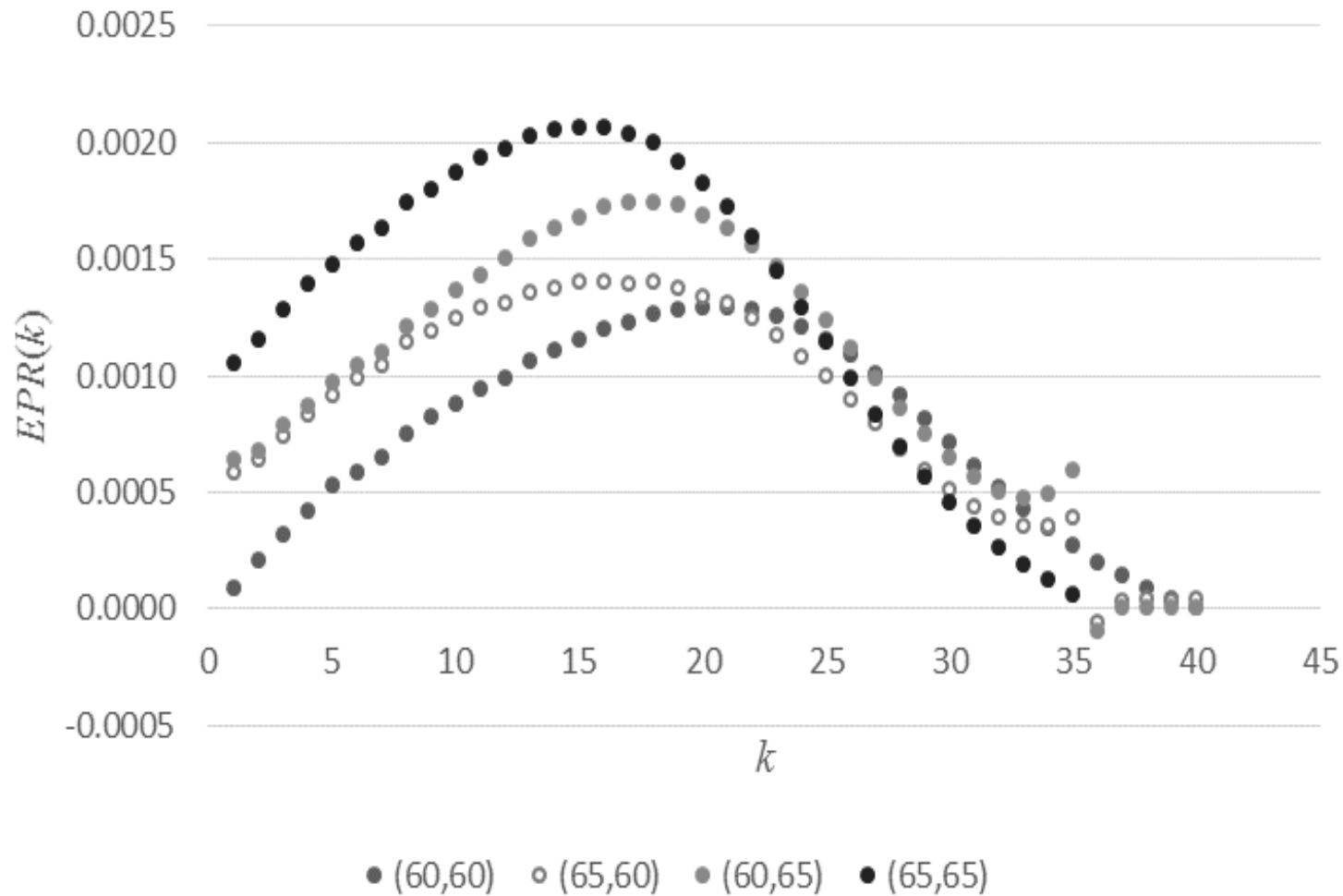
The interest rate determined based on safe financial instruments ($i = 2.1\%$)

SECOND-ORDER BASIS (real)

T_x^M, T_x^W - dependent

The real interest rate ($i = 3.1\%$)

AVERAGE PROFIT EXPECTED TO EMERGE AT THE END OF YEAR k





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